



# Synchronization between Keyboard Typing and Neural Oscillations

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## Abstract

■ Rhythmic neural activity synchronizes with certain rhythmic behaviors, such as breathing, sniffing, saccades, and speech. The extent to which neural oscillations synchronize with higher-level and more complex behaviors is largely unknown. Here, we investigated electrophysiological synchronization with keyboard typing, which is an omnipresent behavior daily engaged by an uncountably large number of people. Keyboard typing is rhythmic, with frequency characteristics roughly the same as neural oscillatory dynamics associated with cognitive control, notably through midfrontal theta (4–7 Hz) oscillations. We tested the hypothesis that synchronization occurs between typing and midfrontal theta and breaks down when errors are committed. Thirty healthy participants typed words and sentences on a keyboard without visual feedback, while EEG was recorded. Typing rhythmicity was investigated by interkeystroke interval analyses and by

a kernel density estimation method. We used a multivariate spatial filtering technique to investigate frequency-specific synchronization between typing and neuronal oscillations. Our results demonstrate theta rhythmicity in typing (around 6.5 Hz) through the two different behavioral analyses. Synchronization between typing and neuronal oscillations occurred at frequencies ranging from 4 to 15 Hz, but to a larger extent for lower frequencies. However, peak synchronization frequency was idiosyncratic across participants, therefore not specific to theta nor to midfrontal regions, and correlated somewhat with peak typing frequency. Errors and trials associated with stronger cognitive control were not associated with changes in synchronization at any frequency. As a whole, this study shows that brain-behavior synchronization does occur during keyboard typing but is not specific to midfrontal theta. ■

## INTRODUCTION

In the past few decades, some sensory processes and behaviors have been shown to be rhythmic. For instance, visual processing oscillates between low and high perception stages (VanRullen, 2016), and behaviors such as speech (Aiken & Picton, 2008), although not strictly rhythmic per se, have also been associated with periodic temporal properties (Rosen, 1992). The phase of rhythmic neural activity (i.e., neural oscillations) has been linked to sensory and motor processes (Drewes & VanRullen, 2011; Gross et al., 2001), and some studies have reported that neural activity can become temporally aligned with rhythmic behaviors (Haegens & Zion Golumbic, 2018; Kösem et al., 2018; Zoefel & VanRullen, 2016; Calderone, Lakatos, Butler, & Castellanos, 2014). On the other hand, neural activity can also provide temporal constraints to behaviors and several cognitive functions have been proposed to be inherently rhythmic in the sense that they are organized according to the timing of neural oscillations. Notably, this has been shown for attention (Fiebelkorn & Kastner, 2019) and cognitive control (Duprez, Gulbinaite, & Cohen, 2020).

Most of the evidence of synchronization between behavior or cognitive functioning with brain oscillatory activity have been inferred from simple experimental psychology tasks. Although the studies using such paradigms are very informative, they lack the complexity of real-life behaviors. Thus, it is important to understand whether neural oscillations and behavior synchronization extend to more naturalistic behavior. However, such naturalistic behavior should also comply with the same experimental rigor required in usual tasks, leading to difficulties in finding a good compromise between rigor and behavioral relevance.

Fortunately, technological advances also led to the emergence of newly common human behaviors. A good example of such recent and widespread behavior is keyboard typing, which is omnipresent in our modern societies, whether it is expressed on personal computers, laptops, smartphones, or tablets. Studying this behavior has several advantages such as an ease in laboratory use (simultaneous recording of behavior and EEG with every stimulus and keypress as an event marker) and in finding performant typists. Furthermore, some evidence suggests that typing is also a rhythmic behavior with mean interkeystroke intervals (IKIs) of approximately 135 msec corresponding to a 7-Hz typing frequency, which is in the theta range (Yamaguchi, Crump, & Logan, 2013). More importantly, regarding brain activity, recent results suggest

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that typing has similar neural substrates to cognitive control and error monitoring. Indeed, typing errors are associated with stronger midline frontal theta power (Kalfaoğlu, Stafford, & Milne, 2018), which is commonly reported in more standard cognitive control tasks (Cavanagh & Frank, 2014; Cohen & van Gaal, 2014).

In this study, our goal was to test the hypothesis that keystrokes during typing behavior would occur at consistent midfrontal theta oscillatory phase, which would indicate synchronization between keyboard typing and midfrontal theta activity. We also expected that typing errors would be associated with a decrease in phase consistency over keystrokes, thus showing a decrease in brain–behavior synchronization during typing. To this end, we used a typing task that required participants to type both real words/sentences, or pseudowords/sentences, in which case the need for cognitive control would be greater. We first focused on showing rhythmicity in typing and then used frequency-specific spatial filters based on a multivariate guided source separation method (generalized eigen decomposition [GED]; Cohen, 2017) to investigate synchronization between typing and neuronal oscillations.

## METHODS

### Participants

Thirty healthy native Dutch speakers (16 women, 14 men, mean age = 22 years old,  $SD = 3.12$  years) were recruited for this study through the Sona recruitment system of the Donders Institute for Brain, Cognition and Behavior and were compensated with €20 or course credit. Selected participants self-reported being able to touch-type. Twenty-six participants were right-handed, and eight participants reported using four to eight fingers instead of 10 for typing. We chose not to exclude left-handers because typing requires both hands, and we wanted to sample from the general population. All participants signed written informed consent, and this study was approved by the Donders Institutes' ethics committee. Following the preprocessing steps detailed below, three participants were excluded from the analyses.

### Task

The typing task was based on the one used in Kalfaoğlu and Stafford (2014). Participants were seated in front of a 24-in. IRT monitor and had to type the presented stimuli on a Dell qwerty-keyboard as quickly and as accurately as possible. In different blocks, we randomly presented real words, real sentences, pseudowords, or pseudosentences. Participants were instructed to type what they saw without looking at the keyboard and to press the enter key when finished. An intertrial interval of 2 sec separated subsequent trials. Each block lasted 5 min. There were also five “free typing” trials during which participants responded to open questions about their activities the day before, of

book/film plots, for 3 min with no other restrictions than typing correct Dutch sentences. These trials were not included in the analyses. No visual feedback was provided to the participants during typing. Participants were observed via a video camera and verbally instructed not to look at the keyboard when needed. The task was programmed in MATLAB using functions from Psychtoolbox (Brainard, 1997) to integrate the stimuli with the hardware and to send triggers for each keystroke to the EEG acquisition, with high temporal precision.

Participants were instructed to correct typing errors by pressing the backspace key until they recall being at the position of the error. They were also informed that punctuation and capital letters could be ignored. Throughout the task, the average typing speed of the participants was calculated and used as a threshold. Should they type below the threshold, the participants would be asked to type faster (except in the free typing condition) by a “Type faster” message followed by a 1-sec intertrial interval. Stimuli consisted of 200 real words (average of 10.35 characters,  $SD = 2.5$ ), 200 pseudowords (average of 9.0 characters,  $SD = 1.8$ ), 100 real sentences (average of 60.8 characters including spaces,  $SD = 13.6$ ), and 100 pseudosentences (average of 59.6 characters including spaces,  $SD = 9.6$ ). Pseudowords were taken from a pseudoword generator (Wuggy; Keuleers & Brysbaert, 2010), and sentences were taken from Drijvers, Mulder, and Ernestus (2016). Pseudosentences were created by replacing words with pseudowords (from the same generator) in the sentences. Given that we had no independent typing task, we estimated participants' typing abilities based on the sentence condition. Typing speed was calculated as the number of words per minute. Participants had an average of 31.8 wpm ( $SD = 6.5$ ) when speed was calculated on five-letter words and higher, 46.6 wpm ( $SD = 8.9$ ) when calculated on four-letter words and higher, and 74.6 wpm ( $SD = 14.2$ ) when calculated on all words. Overall, 79.9 % ( $SD = 7.3$ ) of words contained no error in the sentence condition and there was 94.6 % ( $SD = 1.9$ ) correctly typed characters.

### EEG Acquisition and Preprocessing

EEG data were acquired at 1000 Hz using a cap with 64 active electrodes placed according to the “M10” equidistant electrode layout. Four additional electrodes were used to record horizontal and vertical eye movements. The EEG signal was amplified by BrainAmp amplifiers driven by a BrainVision Powerpack (Brain Products GmbH). The signal was recorded using the Brain Vision software. Keystrokes were recorded as event markers in the EEG signal by sending transistor–transistor logic pulses. For each participant, electrode locations were recorded as Polhemus  $x$ - $y$ - $z$  coordinates and were further imported in EEGLAB (Delorme & Makeig, 2004), which was used for all further preprocessing steps. EEG data were filtered between 0.5 and 40 Hz. Continuous EEG was visually

inspected to identify bad channels, which were removed and interpolated. Three participants were excluded because of excessive EEG artifacts, leaving 27 participants for further analyses. After these steps, EEG was rereferenced to an average reference. Independent component analysis was performed to identify artifacts such as eyeblinks using the *jader* function of EEGLAB (Delorme & Makeig, 2004). Visual inspection of the components' topography, activity, and power spectrum was performed to identify bad components that were then removed from the signal (1.3 components on average were removed). Specific ASCII code marker numbers were used to track important events, such as stimulus display, letter keystrokes, backspace presses, and return button presses.

For all subsequent analyses, trials were subdivided into four conditions: words, sentences, pseudowords, and pseudosentences; and into three trial outcomes: correct trials, corrected errors (which were defined by the occurrence of a backspace press), and other errors, which contained all uncorrected errors that arose from character omission/addition, wrong character typed, missing words in sentences, and so forth. We dissociated corrected from uncorrected errors because we expected behavioral and EEG activity changes associated with error awareness (Kalfaoğlu et al., 2018). Punctuation errors were not included in the analyses. Importantly, for the RT analyses, trials were labeled as an error when they contained at least one error. In the IKI and EEG signal analyses, correct words in error (pseudo)sentences were treated as correct in the analyses. However, as long as a (pseudo)word contained at least one error, the entire word was treated as an error.

## Behavioral Analyses

### *RT and Accuracy*

RT was defined as the time between stimulus presentation and the first keystroke. All RTs faster than 200 msec or slower than 3 *SDs* from the mean were excluded from further analyses. RTs were then compared as a function of the linguistic nature of the stimuli (real or pseudowords/sentences), the type of stimuli (words/sentences), and outcome of the trial (correct, corrected error, other error). All statistical analyses were done using R (Version 3.6.1; R Core Team, 2019). We performed linear mixed-model analyses on participant-averaged Box-Cox-transformed RT using the *lme* function of the *{nlme}* package (Pinheiro, Bates, DebRoy, Sarkar, & R Core Team, 2020) with a 2 (linguistic nature) + 2 (nature of the stimuli) + 3 (trial outcomes) design, with participant as a random intercept. We applied this transform because graphical check of the models' residuals showed better compliance with the models' assumptions when data were transformed than when left raw. We tested log, inverse, and square-root transforms and found better compliance with the Box-Cox transform based on visual inspection of the models' residuals. We

focused the analysis on participant-averaged RT because we wanted a simple model accounting for unexplained differences between participants and because modeling all trials' RT explained less variance as compared to participant-averaged RT (based on the marginal and conditional  $R^2$  values). We chose not to investigate the interaction between linguistic nature, nature of the stimuli, and trial outcome as further interpreting all the different contrasts would prove difficult and would not be further relevant for this study. Accuracy data (Box-Cox transformed) was analyzed following the same statistical design as for RT with a 2 (linguistic nature) + 2 (nature of the stimuli) + 3 (trial outcomes) design, with participant as a random intercept.

Graphical checks of the model's residuals distribution were performed to ensure compliance with the model's assumptions. For all mixed models, marginal and conditional  $R^2$  were calculated using the *{MuMin}* package (Barton, 2009) and are reported in the results. The *anova* function of the *{car}* package (Fox & Weisberg, 2019) was used on the models to extract significance of the fixed effects by the means of *F* tests. In the case of significant main effects, post hoc comparisons were performed using the *glht* function of the *{multcomp}* package that provides adjusted *p* values using individual *z* tests (Hothorn, Bretz, & Westfall, 2008). The significance threshold was set to  $p = .05$ .

### *IKIs*

As a first method to investigate rhythmicity in typing, we focused on IKIs, which corresponds to the time between consecutive keystrokes. For this analysis, we focused on stimuli that contained at least four characters. Average IKI for each condition, outcome, and participant was calculated. Then, comparison between experimental conditions were carried out using similar linear mixed models as described for RT but applied on log-transformed IKI, which showed better compliance with the models' assumptions.

### *Kernel Density Estimation*

We quantified behavioral rhythmicity in typing by the mean of a kernel density estimation (KDE) analysis. This method is a nonparametric calculation of a probability density function. For each word, we computed the sum of the convolution of each IKI with a Gaussian kernel. The Gaussian kernel was defined as follows:

$$e^{-4 \ln(2)t^2/s^2} \quad (1)$$

where  $t$  is a time vector and  $s$  is the kernel bandwidth that was set at 1. At the IKI level, this was done by multiplying the Fourier transform of the kernel with the Fourier transform of the IKI signal and then taking the inverse Fourier transform of the result. The IKI signal is a 0.1-Hz resolution vector ranging from 1 to 100 Hz containing only zeros

except for the closest frequency of the IKI (which contains 1). For each word, the result of KDE is the sum of each IKI convolution normalized by the number of letters composing the word. The average result of KDE was calculated for each participant according to experimental conditions and outcome. Further analyses involved identifying the peak typing frequency for each condition and outcome and comparing them at the group level. We used the same linear mixed models as previously described. In that case, peak frequencies were Box-Cox-transformed before modeling because inspection of the model's residuals indicated better compliance with the model's assumption in that case.

### Spatial Filtering of the Data

All preprocessed EEG signal analyses were performed in MATLAB (The MathWorks, Version 2016a) using custom-written code (Cohen, 2014). In order to investigate frequency-specific synchronization between brain activity and typing behavior, we applied spatial filters constructed via GED: For each frequency, a spatial filter was used to extract a single component that maximizes activity at that specific frequency. This spatial filter maximizes the separation between the target frequency and broadband activity with a specific topographical distribution. Critically, the spatial filter does not eliminate other frequencies activity: It selects a set of channel weights that identifies the maximum energy at the target frequency on the scalp. The broadband data are then passed through this spatial filter, and a time–frequency analysis can be applied to that component time series. Designing the spatial filter involves computing channel covariance matrices and finding a set of channel weights that maximally differentiates the narrowband covariance from the broadband covariance (de Cheveigné & Arzounian, 2015; Nikulin, Nolte, & Curio, 2011). Calculation is done via Rayleigh's quotient, finding a set of channel weights in vector  $\mathbf{w}$  that maximizes  $\lambda$ :

$$\lambda = \frac{\mathbf{w}^T \mathbf{S} \mathbf{w}}{\mathbf{w}^T \mathbf{R} \mathbf{w}} \quad (2)$$

with  $\mathbf{S}$  as the covariance matrix of the signal narrow-band filtered at the studied frequency, and  $\mathbf{R}$  the covariance matrix from the broadband signal.

Equation 2 can be extended as a generalized eigenvalue equation:

$$\mathbf{S} \mathbf{W} = \mathbf{R} \mathbf{W} \Lambda \quad (3)$$

The column of  $\mathbf{W}$  with the highest associated eigenvalue in  $\Lambda$  corresponds to the spatial filter that maximally distinguishes  $\mathbf{S}$  from  $\mathbf{R}$ . The spatial filter can then be applied to the data as follows:

$$\mathbf{y} = \mathbf{w}^T \mathbf{X} \quad (4)$$

with  $\mathbf{w}$  being the column of  $\mathbf{W}$  corresponding to the largest associated eigenvalue, and with  $\mathbf{X}$  the Channel  $\times$

Time data matrix. The resulting component time series is then the row vector  $\mathbf{y}$  on which all subsequent analyses can be performed. This spatial filter method has the advantage of increasing signal-to-noise ratio, taking into account interindividual topographical differences, and avoiding electrode selection bias (Cohen, 2017; Nikulin et al., 2011).

In this study, we investigated frequencies from 3 to 15 Hz with a 0.5-Hz resolution. For each GED, narrow-band filtering of the data to create the  $\mathbf{S}$  matrix was done using these frequencies. We limited the studied frequency to 15 Hz because we can reasonably assume that it is unlikely that one could type faster than 15-Hz frequency.

The selection of the best spatial filter  $\mathbf{w}$  in the  $\mathbf{W}$  result matrix was done through a combination of a statistical criterion (the size of the eigenvalues) and visual inspection of the topographical maps of the spatial filters' activation pattern. We inspected the 15 spatial filters associated with 15 highest eigenvalues and then removed the filters that were clearly artifactual (for instance, showing clear eye-blink activity, or one/several electrode issues). After these steps, the component with the highest eigenvalue was chosen as the spatial filter regardless of the topographical activation pattern.

Finally, one should note that GED spatial filtering does not allow us to make inferences at the cortical level and all further analyses and interpretations are limited to topographical, channel-level distributions.

### Time–Frequency Decomposition

After spatial filtering of the data, time–frequency decomposition was performed to inspect time–frequency power of the components defined by GED. To this end, we used complex Morlet wavelet convolution: The Fourier transform of the data (GED component) was multiplied by the Fourier transform of a set of complex Morlet wavelet and then taking the inverse Fourier transform of the result. Complex Morlet wavelets were defined as follows:

$$e^{i2\pi ft} e^{-t^2/2s^2} \quad (5)$$

where  $t$  is time,  $f$  is frequency ranging from 1 to 50 Hz (in 60 logarithmically spaced steps), and  $s$  is the width of each frequency band, defined as  $n/(2\pi f)$ , with  $n$  increasing logarithmically from 4 to 12. We then took the squared-magnitude of the resulting signal ( $Z$ ) as *real*  $[Z(t)^2] + \textit{imaginary} [Z(t)^2]$  to obtain power at each frequency. Power was then baseline-corrected using a decibel (dB) transform:  $\textit{dB power} = 10 \times \log_{10}(\textit{power} / \textit{baseline})$ . Time–frequency decompositions were applied either on the data epoched from  $-700$  to  $1500$  msec around stimulus presentation. Baseline power was defined as the average power across all conditions in the period ranging from  $-500$  to  $-200$  msec before stimulus onset. When plotting time–frequency power results, we deliberately used a different colormap than for the topographical maps



of the spatial filters activation (see Figure 5C–5F and 5A–5D) in order to differentiate between power and activation of the spatial filter, which are completely different measures. Significant power changes from baseline were investigated by using permutation analyses. The null hypothesis was constructed by cutting the time–frequency map of power at a random time point and placing the second part before the first part, thus misaligning the time of stimulus onset. This operation was carried out 500 times and resulted in a distribution of time–frequency maps under the null hypothesis. Then, the real time–frequency map was *z*-scored by subtracting the average power value under the null hypothesis from real power and dividing it by the standard deviation. The significance threshold was set at  $p = .05$ .

### Brain–Behavior Synchronization Measures

We tested for brain–keystroke synchronization by extracting the phase angle time series via the filter–Hilbert transform method. We computed the consistency of brain phase angles extracted from the time points of the keystrokes over all trials in a condition-specific manner. Regarding errors, we extracted the phases of all the keystrokes in erroneous words in the (pseudo)word and (pseudo)sentence conditions. The phases of the keystrokes of correctly typed words in error (pseudo)sentences were not used to compute error phase synchronization. Phase consistency was computed as follows:

$$w = \left| n^{-1} \sum_{r=1}^n e^{ik_r} \right| \quad (6)$$

$n$  is the number of keystrokes,  $e^{ik}$  is the complex representation of the phase angle  $k$  of keystroke  $r$ . Values close to 0 indicate random distribution of the phase angles, whereas values close to 1 indicate strong phase consistency.

At the participant level, the significance of phase consistency was calculated using the following equation:

$$p = e^{-Nc^2} \quad (7)$$

with  $N$  being the number of keystrokes and  $c$  the consistency as calculated using Equation 6.

To compare phase consistency between frequency, condition, and outcome, we applied a *z*-score normalization based on permutation analyses to the data. At each iteration, a phase angle time series vector was created using a cut-and-swap procedure that cuts the original time series at a random point and puts the second part before the first part, thus misaligning keystrokes and phases' timing. Following this, phase consistency was calculated as described in Equation 4. This procedure was carried out 500 times separately for each condition and outcome, providing a distribution of phase consistency under the null hypothesis that phase angles are randomly distributed. *Z*-score normalization was then applied to real consistency values by subtracting the average consistency of the null

hypothesis distribution and dividing the result by its standard deviation.

To investigate frequency-specific effects of outcome and experimental condition on phase consistency, linear mixed models with a 2 (linguistic nature) + 2 (nature of the stimuli) + 3 (outcomes) design, and with participant as a random effect, were performed at each frequency on box-cox transformed consistency values (using the same functions as described in the RT and Accuracy section).

### Data and Code Availability

All EEG and behavioral data are posted on the Donders data repository. All the codes used for EEG analyses, statistical analyses, and figures are openly available at [github.com/jduprez](https://github.com/jduprez).

## RESULTS

### Behavioral Results—Is There Rhythmicity in Typing?

#### RT and Accuracy

As shown in Table 1, accuracy depended on the type of stimulus, with overall lower accuracy for (pseudo)sentences than (pseudo)words,  $F(1, 70) = 394.5, p < .0001$ ; conditional  $R^2 = .84$ , marginal  $R^2 = .67$ . However, whether the stimulus had a linguistic nature or not did not influence accuracy,  $F(1, 70) = 1.5, p = .22$ . The low accuracy for (pseudo)sentences is explained by the fact that the entire sentence was labeled as an error if at least one word had an error. Indeed, sentences and pseudosentences had respectively an average of 2.3 and 2.2 errors per trial when erroneous (Table 1). Figure 1 displays the group-average and distribution of RT showing that RT increases when the stimuli was either nonlinguistic or a sentence. This pattern was confirmed by the statistical analyses with a significant Linguistic Nature effect,  $F(1, 293) = 19, p < .0001$ ; conditional  $R^2 = .9$ , marginal  $R^2 = .2$ , as well as a significant Type of Stimulus effect,  $F(1, 293) = 645, p < .0001$ . Real stimuli had lower RT (795.4 msec,  $SD = 132.6$ ) than pseudostimuli (825.4 msec,  $SD = 170.7$ ), and words had shorter RT (749.7 msec,  $SD = 134.3$ ) than (pseudo)sentences (871.1 msec,  $SD = 147.4$ ). Although graphical interpretation would suggest that corrected and other errors had shorter RT than correct trials, this was not supported by the statistical analyses, which showed no significant Trial Outcome effect,  $F(2, 292) = 0.07, p = .933$ . As a whole, these data show an increase in RT between nonlinguistic stimuli and (pseudo)sentences as compared to linguistic stimuli and (pseudo)words.

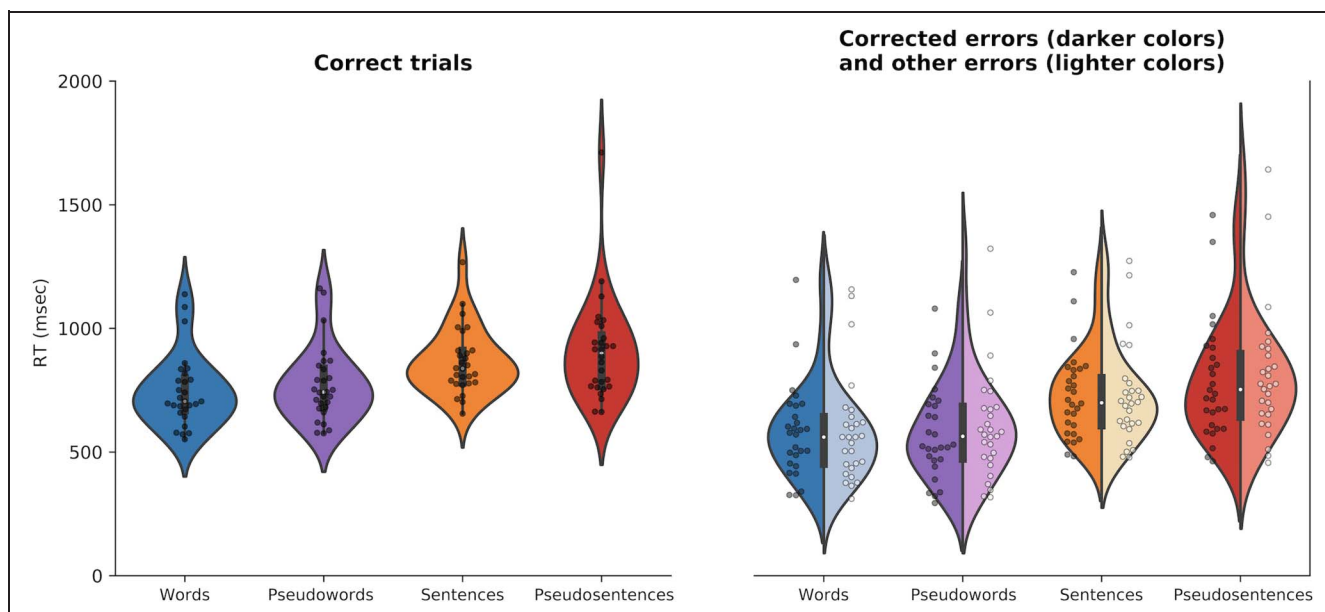
#### IKIs

Overall IKI had an average of 156.2 msec ( $SD = 52.6$ ), corresponding to 6.4 Hz. Figure 2 shows that IKIs changed according to condition. An increase in IKI (thus a decrease

**Table 1.** Average Number of Trials and Accuracy According to Experimental Conditions and Trial Outcome

	<i>Number of Correct Trials</i>	<i>% Correct Trials</i>	<i>Number of Corrected Errors</i>	<i>% Corrected Errors</i>	<i>Number of Other Errors</i>	<i>% Other Errors</i>	<i>n of Erroneous Words</i>	<i>% of Correct Words in Sentences</i>
Pseudowords	72 ( <i>SD</i> = 18)	68.5 ( <i>SD</i> = 12.2)	13 ( <i>SD</i> = 5)	13.2 ( <i>SD</i> = 5.6)	19 ( <i>SD</i> = 11)	18.3 ( <i>SD</i> = 10.5)		
Pseudosentences	9 ( <i>SD</i> = 6)	20.2 ( <i>SD</i> = 17.1)	26 ( <i>SD</i> = 8)	49.8 ( <i>SD</i> = 15.1)	17 ( <i>SD</i> = 8)	32.5 ( <i>SD</i> = 14.3)	2.2 ( <i>SD</i> = 0.7; range: 0.1–3.6)	75.7 ( <i>SD</i> = 10.4)
Words	71 ( <i>SD</i> = 20)	64.8 ( <i>SD</i> = 13)	13 ( <i>SD</i> = 6)	13.3 ( <i>SD</i> = 7.6)	21 ( <i>SD</i> = 12)	20 ( <i>SD</i> = 10.9)		
Sentences	14 ( <i>SD</i> = 8)	27.9 ( <i>SD</i> = 17.1)	21 ( <i>SD</i> = 5)	41 ( <i>SD</i> = 11.5)	18 ( <i>SD</i> = 8)	33.4 ( <i>SD</i> = 12.2)	2.3 ( <i>SD</i> = 0.8; range: 0.2–3.7)	70.8 ( <i>SD</i> = 10.5)
Average number of backspaces per word in corrected errors	1.5 ( <i>SD</i> = 0.2)							
Overall average number of backspaces	115.9 ( <i>SD</i> = 36.3)							
Average backspace IKI	383.3 ( <i>SD</i> = 48.7)							

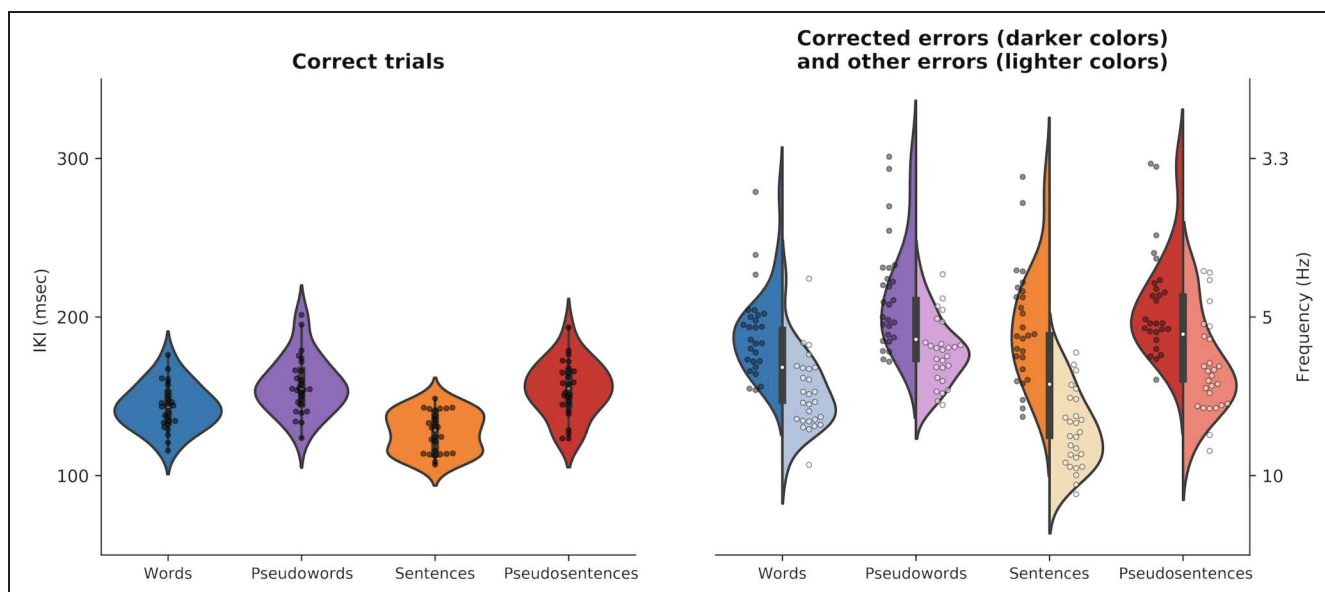
The lower part of this table contains backspace data details.



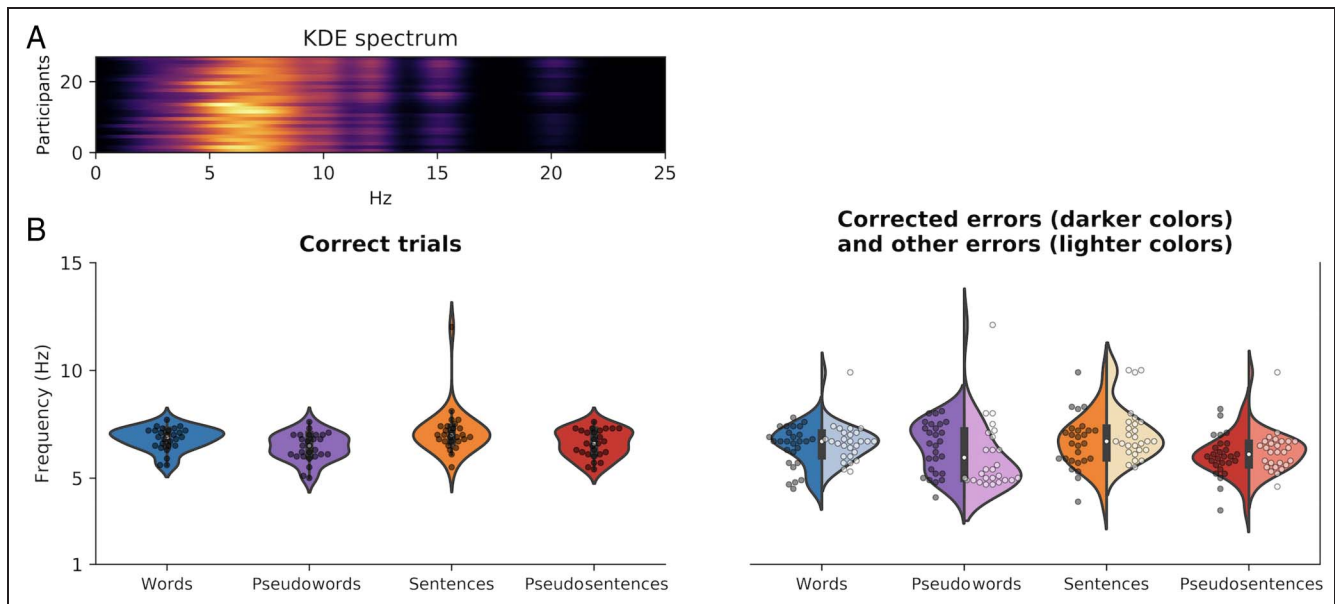
**Figure 1.** RT of the first keystroke after stimulus display according to condition and trial outcome. Violin plots show the distribution of the data with an inset boxplot. Each data point corresponds to the average RT of a participant. The right part of the figure shows asymmetrical violins that present both corrected errors data (in darker colors) and other errors data (in lighter colors).

in typing frequency) can be observed nonlinguistic (178.2 msec,  $SD = 30.1$ ) compared to linguistic stimuli (157.3 msec,  $SD = 31.5$ ) and for (pseudo)words (172.1 msec,  $SD = 29$ ) compared to (pseudo)sentences (164.4 msec,  $SD = 34.8$ ). These differences were supported by the statistical analyses with a significant Linguistic Nature effect,  $F(2, 282) = 225.6, p < .0001$ ; conditional  $R^2 = .83$ , marginal  $R^2 = .58$ , and a significant Type of Stimulus effect,  $F(2, 282) = 38.04, p < .0001$ . Figure 3 also suggests that IKI varied between trial outcomes with shorter IKI for correct

trials than for corrected and other errors. We found a significant Trial Outcome effect,  $F(2, 282) = 407.95, p < .0001$ , and post hoc analyses confirmed that correct trials had shorter IKI (145.4 msec,  $SD = 19.5$ ) than other errors (161 msec,  $SD = 24.5$ ), which had shorter IKI than corrected errors (196.8 msec,  $SD = 27$ ; all  $p < .0001$ ). However, given that backspace IKI were greater than other IKIs (383.3 msec,  $SD = 48.7$ ; Table 1), we reran the analysis to check whether the difference between corrected errors and other outcomes was not driven by backspace IKIs.



**Figure 2.** IKI according to condition and trial outcome. Violin plots show the distribution of the data with an inset boxplot. Each data point corresponds to the average IKI of a participant. The right part of the figure shows asymmetrical violins that present both corrected errors data (in darker colors) and other errors data (in lighter colors). A second y-axis has been added on the right to indicate the typing frequency corresponding to the IKI.



**Figure 3.** (A) Participant-specific spectrum of KDE results (averaged over all conditions for each participant - grand average of all stimuli IKI probability density function); (B) Peak frequency extracted from the KDE result according to condition and trial outcome. Violin plots show the distribution of the data with an inset boxplot. Each data point corresponds to the peak typing frequency of a participant. The right part of the figure shows asymmetrical violins that present both corrected errors data (in darker colors) and other errors data (in lighter colors).

When excluding backspace IKIs, the Outcome effect remained,  $F(2, 282) = 84.71, p < .0001$ , although corrected errors had a shorter IKI (167.5 msec,  $SD = 25.2$ ) than in the previous analysis (196.8 msec,  $SD = 27$ ). The post hoc comparison showed that the difference between correct errors and other errors remained significant as well ( $p = .0006$ ). As a whole, typing frequency was faster for real than pseudostimuli and for (pseudo)words than (pseudo)sentences stimuli. It is worth noting that typing frequency was also faster for other errors than for corrected errors.

#### KDE Analyses

Using KDE on IKI data provided similar results as raw IKI analyses regarding typing frequency. Inspection of the KDE spectrum (Figure 3A) shows that typing mostly occurs around 6 Hz. In order to better assess typing frequency at the group level, we extracted the peak frequency for each participant according to condition and trial outcome (Figure 3B). In that case, results were less clear than when focusing on IKI. Indeed, although graphical inspection seems to show similar typing frequency regardless of condition and trial outcome, statistical analyses revealed significant differences for the Linguistic Nature effect (Linguistic Nature:  $F(1, 293) = 37.9, p < .0001$ ; conditional  $R^2 = .45$ , marginal  $R^2 = .10$ ) with higher frequencies for real stimuli (6.8 Hz,  $SD = 1$ ) than pseudostimuli (6.3 Hz,  $SD = 1.1$ ). We also found a significant outcome effect,  $F(2, 293) = 9.02, p = .0002$ , but the type of stimulus did not significantly influenced frequency,  $F(1, 293) = 3.4, p = .064$ . It is worth noting that, in this analysis, conditional and marginal  $R^2$  were lower than for RT or IKI, indicating less variance explained by the

model. When applying post hoc tests, we found that correct trials had faster frequency (6.8 Hz,  $SD = 0.8$ ) than corrected errors (6.4 Hz,  $SD = 0.9; p < .001$ ) and other errors (6.5 Hz,  $SD = 1.3; p = .003$ ). The comparison between corrected errors and other errors was not significant ( $p = .74$ ). Overall average peak frequency was at 6.5 Hz ( $SD = 1.1$ ), which is consistent with the typing frequency obtained in the IKI results (6.4 Hz).

#### EEG—Does Typing Synchronize with Brain Oscillatory Activity?

##### Topographies of the Spatial Filters

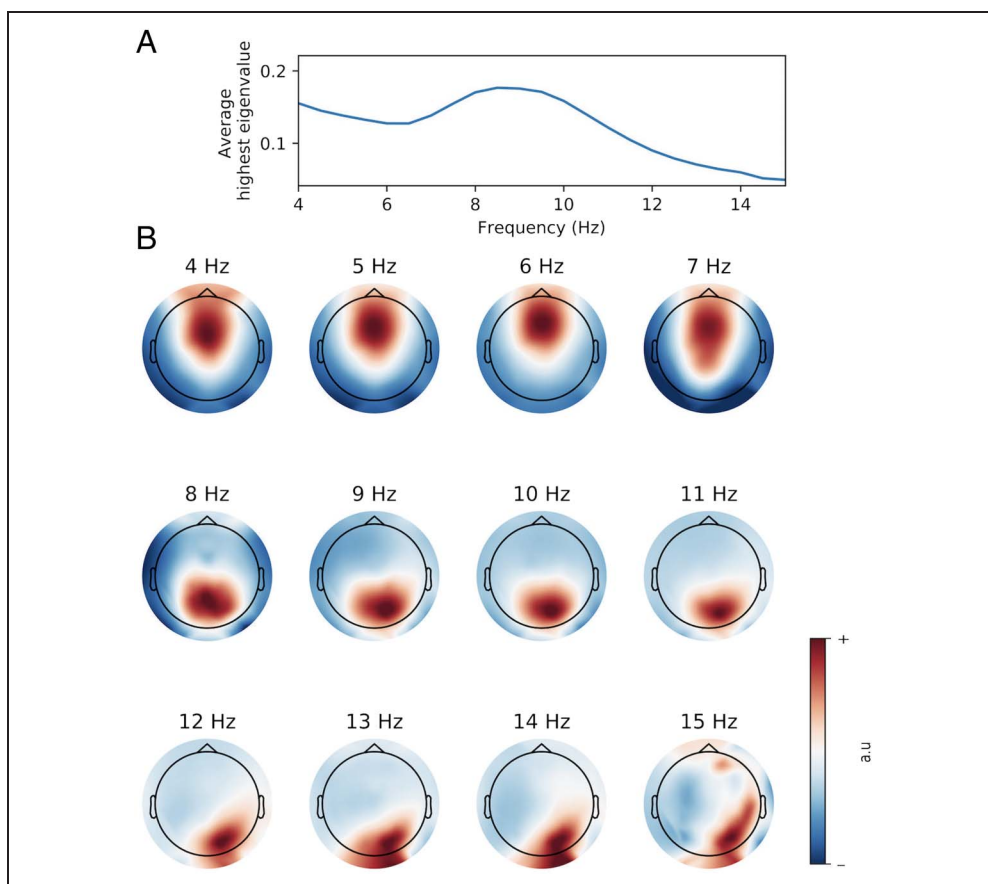
Although we investigated frequencies ranging from 3 to 15 Hz, we only report EEG results from 4 to 15 Hz because 3-Hz activity was dominated by eye-level activity and because no participant typed at 3 Hz. Figure 4A shows the size of the largest eigenvalue according to frequencies at the group level. The highest values were observed at 8.5 Hz, which indicates that this frequency was more easily differentiated from broadband activity than other frequencies, especially from frequencies above 12 Hz. Inspection of the spatial filters created by GED revealed clear activation patterns (Figure 4B). Greater activation was centered around midfrontal electrodes from 4 to 7 Hz. A shift from midfrontal to occipital activation occurred from 8 to 11 Hz, which became more occipito-parietal from 12 to 15 Hz.

##### ERP and Time-Frequency Analyses

In order to further evaluate the quality of the components created by GED, we inspected the ERP and time-frequency decompositions of the components. Here, we



**Figure 4.** (A) Spectrum of group-averaged highest eigenvalue. (B) Topographical maps showing the activation pattern (in arbitrary units) of the spatial filters designed by GED for each frequency. Although the spatial filters were computed at 0.5-Hz resolution, only 1-Hz resolution topographical maps are shown here.



report the results of two frequency components: the 6-Hz component, because of our hypothesis on theta synchronization with typing, and the 8.5-Hz component (Figure 5) given that it is the one associated with the highest eigenvalues.

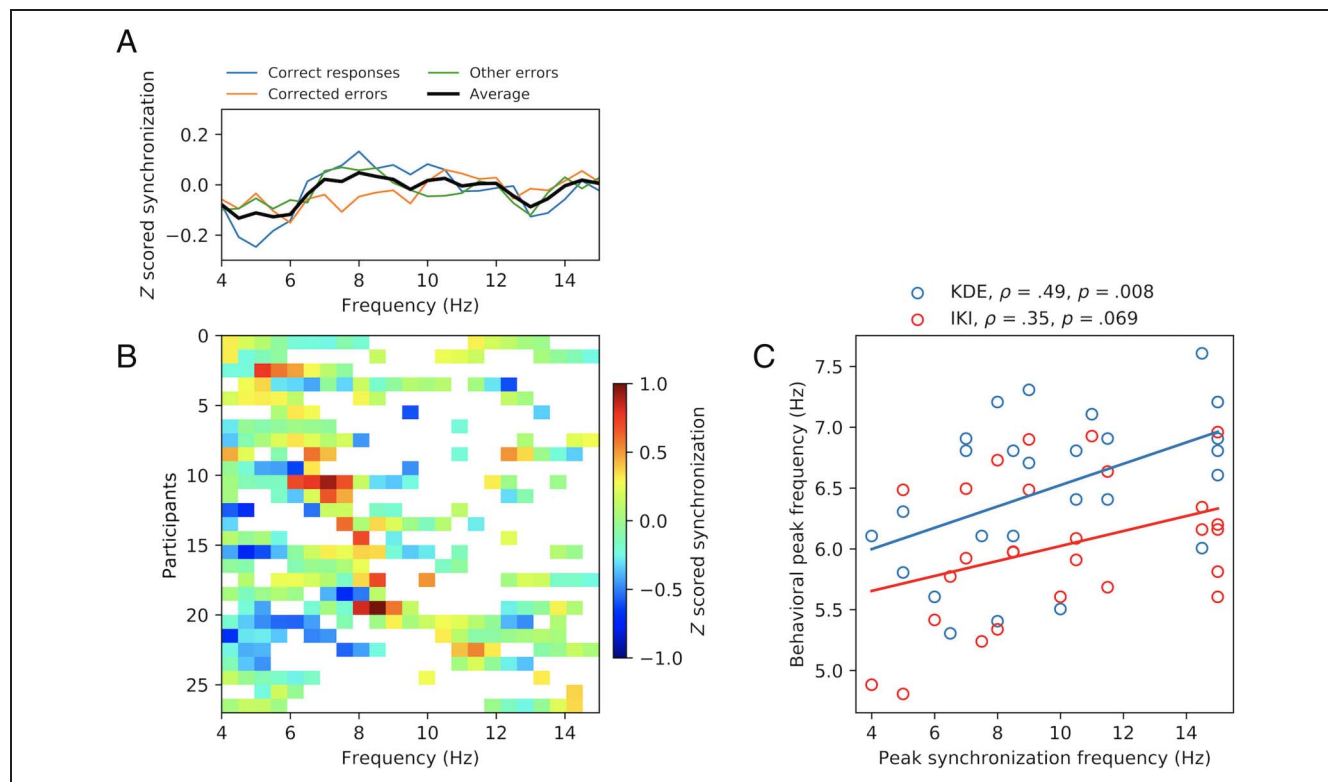
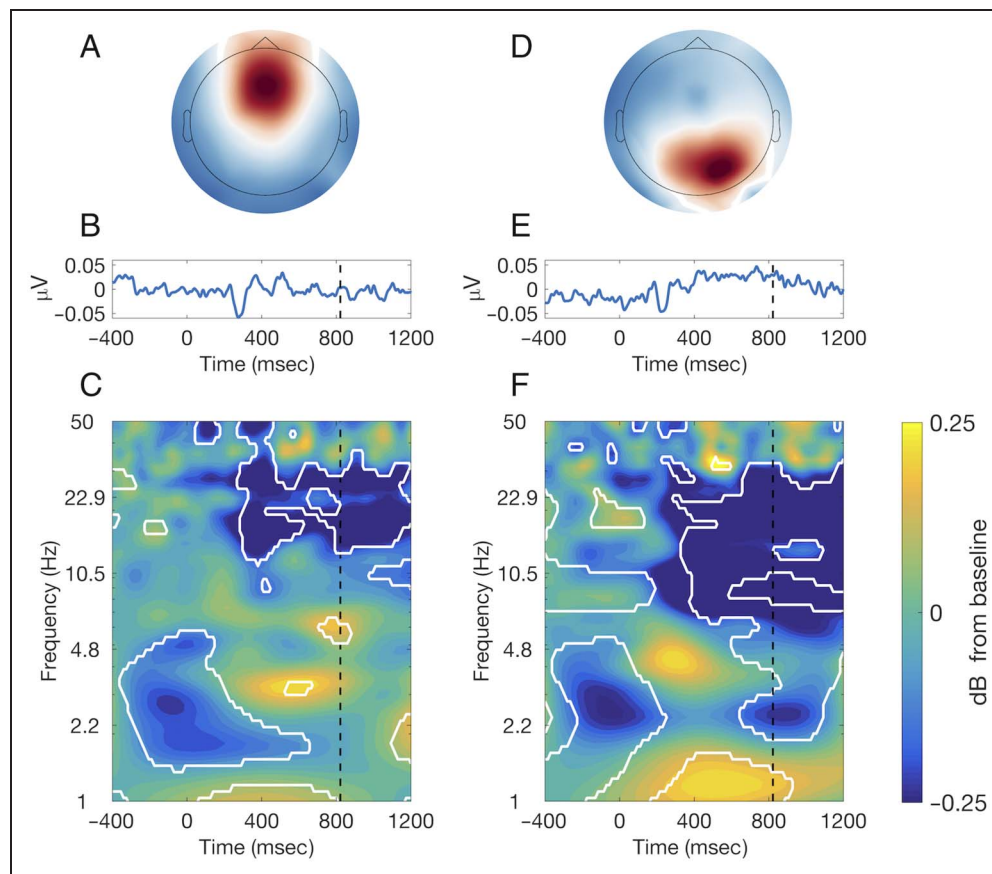
Analyzing the 6- and 8.5-Hz components revealed that the spatial filtering designed by GED resulted in physiological signals that have a clear topography, as well as time-resolved and time-frequency characteristics. Figure 5B and 5E show the group-level ERP with clear positive and negative variations after stimulus onset. Figure 5C and 5F reveal the time-frequency power dynamics. A significant increase in power can be observed around 600 msec after stimulus presentation for the 6-Hz component (thus before mean RT) in the theta band. An important significant decrease in power in the beta band for the 6-Hz component, and that spans the alpha and beta band for the 8.5-Hz component, was also present from roughly 300 to 1200 msec. Very low- and high-frequency changes can also be observed. The topographical map and ERP are very different from one target frequency to the other. Time-frequency power dynamics, although sharing overall similar features, also show different characteristics such as stronger alpha decrease and delta increase for the 8.5-Hz component. As a whole, GED allowed to isolate frequency-specific components that each had its particular activation pattern and oscillatory characteristics.

#### *Frequency-specific Phase Clustering of All Keystrokes*

Our main hypothesis was that typing synchronizes with brain activity around the typing frequency. More specifically, we expected that typing in theta frequency would synchronize with midfrontal theta oscillations. For frequencies ranging from 4 to 15 Hz, phase-behavior synchronization was calculated for the corresponding GED component. Z-score normalization based on permutation analyses was then performed in order to ensure frequencies, experimental conditions, and trial outcomes are comparable.

According to Figure 6A, it seems that average synchronization between typing and brain activity peaked at 8 Hz. However, inspection of the participant-specific spectra of synchronization revealed that the peak synchronization frequency was idiosyncratic across participants and ranged from 4 to 15 Hz (Figure 6B). Since brain-behavior synchronization could vary according to experimental conditions, we applied linear mixed models for each frequency to investigate the effect of the linguistic nature of the stimuli, the type of stimuli, and trial outcome on synchronization. This analysis did not reveal any significant experimental condition, or trial outcome effect at any frequency (all  $p > .05$ ), except for the linguistic nature effect at 13,  $F(1, 326) = 5.2, p = .024$ ; conditional  $R^2 = 0.02$ , marginal  $R^2 = .02$ , and 13.5 Hz,  $F(1, 326) = 6.4, p = .012$ ; conditional  $R^2 = .02$ , marginal  $R^2 = .02$ , with higher consistency for

**Figure 5.** (A) Topographical map showing the activation pattern of the 6-Hz spatial filter; (B) group-averaged event-related potential; (C) group-averaged time–frequency plot of decibel-transformed power (white contour lines indicate significant changes from baseline at  $p < .05$  based on permutation testing). Stimulus onset occurred at time = 0 msec, and the vertical dashed line corresponds to the group-averaged RT. (D–F) Same plots but for the 8.5-Hz spatial filter.



**Figure 6.** (A) Outcome-specific and group-averaged z-scored synchronization according to frequency; (B) participant-specific z-scored synchronization according to frequency sorted according to peak synchronization frequency (upper participants have a lower peak frequency). The white pixels indicate nonsignificant synchronization ( $p > .05$ ) at the participant level (calculated over all trials); all other pixels have participant-level synchronization at  $p < .05$ ; (C) scatterplot showing the relationship between peak synchronization frequency and behavioral peak frequency estimated by the IKI and KDE analyses.

nonlinguistic (13 Hz: 0.02,  $SD = 0.96$ ; 13.5 Hz: 0.06,  $SD = 0.97$ ) than linguistic stimuli (13 Hz:  $-0.2$ ,  $SD = 0.85$ ; 13.5 Hz: 0.18,  $SD = 0.81$ ), and for the type of stimulus effect at 9 Hz with lower consistency for words ( $-0.15$ ,  $SD = 0.98$ ) than sentences ( $-0.1$ ,  $SD = 0.91$ ;  $F(1, 326) = 3.9$ ,  $p = .047$ ; conditional  $R^2 = .03$ , marginal  $R^2 = .01$ ). However, it is important to note that for all frequencies, conditional and marginal  $R^2$  were lower than 0.1, indicating only a small amount of variance was explained by the models. Furthermore, although peak synchronization frequency was highly variable across participants, inspection of synchronization significance at the participant level revealed that most significant synchronizations occurred at frequencies lower than 9 Hz (Figure 6B). Figure 6C displays the relationship between peak synchronization frequency and behavioral peak frequency. If typing and brain activity synchronizes the most at the typing frequency, a strong correlation should appear between typing frequency and the synchronization between typing and brain activity. The correlation between peak frequency estimated through IKI analysis with the peak synchronization was not significant ( $\rho = 0.35$ ,  $p = .069$ ), whereas it was the case for correlation between the KDE-estimated frequency and peak synchronization frequency ( $\rho = 0.49$ ,  $p = .008$ ). Critically, this analysis was not circular given that neural oscillation's phase (extracted at the time of the keystrokes) and IKI/KDE values are independent. As a whole, our results suggest that the brain–behavior synchronization mostly occurs at low frequencies during typing but is more nuanced than originally expected.

## DISCUSSION

The rhythmic nature of brain electrophysiological activity has been shown to be associated with rhythmic behaviors through various mechanisms (synchronization, entrainment, changes in behavioral performances). In this study, we tested the hypothesis that keyboard typing, a real-life behavior that has been suggested to be rhythmic, synchronizes with neuronal oscillations, which would manifest as phase consistency of oscillatory activity over keystrokes during typing behavior. More specifically, we expected that synchronization would preferentially occur with midfrontal theta oscillations, which have been reportedly associated with cognitive control, and recently with keyboard typing. We further hypothesized that phase consistency over keystrokes would decrease during errors. Our results partly support our hypothesis by showing that typing is indeed rhythmic and that it does mostly synchronize with brain activity in low frequencies, although with no strong specificity for midfrontal theta.

### Typing is Rhythmic

Keyboard typing is a highly expressed behavior in our modern societies, which makes it an ecologically valid behavior to investigate. Furthermore, studying typing is facilitated by

the ease in collecting huge amounts of data in a short amount of time and without any special requirements in experimental setup and experimenter training. Typing speed has been the focus of ample research, notably regarding keyboard design and typing performance (Soukoreff & Mackenzie, 1995; Kinkead, 1975). Performances are usually assessed by investigating the number of words per minute, or through the IKI, with some studies suggesting that typing occurs in the theta frequency range (Kalfaoğlu et al., 2018; Kalfaoğlu & Stafford, 2014; Yamaguchi et al., 2013). In this study, we used two different approaches to investigate rhythmicity in typing: either by focusing on raw IKI or by calculating their probability density function using KDE. Both approaches led to results showing that typing rhythmicity occurred in the theta frequency range on average (6.4 and 6.5 Hz for the IKI and KDE analyses, respectively). Our results also suggest that typing RT and typing frequency depended to some extent on the need for cognitive control, on the type of stimulus, and on whether errors were committed. Indeed, pseudo-stimuli were associated with longer RTs and lower typing frequency and sentences were associated with longer RTs and higher typing frequency. Regarding trial outcome, lower typing frequency was observed for errors compared to correct trials. Corrected errors showed lower typing frequency compared to other errors, which could be explained by behavioral adjustments such as posterror slowing because of error awareness and monitoring (Schroder & Moser, 2014; Ullsperger, Fischer, Nigbur, & Endrass, 2014). However, this interpretation is mitigated by the fact that overall backspace IKIs were greater than other IKIs. Removing backspace IKIs from the analyses decreased corrected errors IKIs, indicating that backspaces accounted for at least some part of lower typing frequency associated with corrected errors. Nonetheless, removal of these IKIs still resulted in lower typing frequency in corrected errors compared to other errors. Thus, error monitoring effects might also contribute to this effect. For instance, monitoring processes might have been even more elicited given the absence of visual feedback (Pinet & Nozari, 2020). Following this, the absence of visual feedback might account for lower correction rates compared to other studies (Pinet & Nozari, 2021; Kalfaoğlu et al., 2018). Finally, it is important to note that these results were stronger in the IKI analysis. This could be, at least partly, explained by the fact that IKI analyses were based on participant-averaged IKI, whereas KDE analyses were based on participants' peak typing frequency, which could arguably be more consistent across conditions than average IKI. Nonetheless, the lack of the same effects for the KDE analysis leads to caution in interpreting experimental condition and outcome effects on typing frequency and demands further replications.

### Unmixing Frequency-specific Signals

In this study, we hypothesized a synchronization between typing and midfrontal theta oscillations. In order to

increase the signal-to-noise ratio, to avoid bias in electrode selection and to account for variability in participant's topography, we chose to apply a multivariate guided source separation method, namely, GED (de Cheveigné & Arzounian, 2015; Nikulin et al., 2011). This method allows us to create a single component that best reflects target features of the signal (in this case, a narrow-band frequency-specific signal) and has been shown to be very helpful in maximizing low-frequency features in the EEG signal (Cohen, 2017). Our results showed that this method successfully isolated different narrow-band frequency-specific components. Indeed, theta frequency components showed a midfrontal-centered activation topography, which is highly similar to topographies observed during cognitive control tasks, and were associated with comparable time–frequency dynamics such as preresponse theta power bursts (Duprez et al., 2020; Pastötter, Dreisbach, & Bäuml, 2013; Cohen & Cavanagh, 2011; Nigbur, Ivanova, & Stürmer, 2011). Focusing on higher frequencies, topographies shifted to more posterior, occipital, and occipitoparietal activations and components showed distinct time–frequency dynamics with strong alpha suppression. These topographies and power dynamics are usually associated with visual attention (Zhang, Zhang, Cai, Luo, & Fang, 2019; Bauer, Stenner, Friston, & Dolan, 2014; Fries, Womelsdorf, Oostenveld, & Desimone, 2008) and have already been reported during typing (Scaltritti, Suitner, & Peressotti, 2020).

### Brain–Behavior Synchronization during Typing

Synchronization between brain activity and sensory information or behavior has been repeatedly observed in various situations. Brain oscillations can become entrained and phase-aligned to (quasi)rhythmic stimuli with consequences on cognitive performances (see Hanslmayr, Axmacher, & Inman, 2019, for a review on memory), or they can provide a more or less constrained time frame for specific cognitive processes or behavior to occur (Duprez et al., 2020; Fiebelkorn & Kastner, 2019; Miller, Lundqvist, & Bastos, 2018). The similarities between keyboard typing and cognitive control in terms of topographical theta activity and time–frequency power (Kalfaoglu et al., 2018; Cavanagh & Frank, 2014) led us to hypothesize the existence of synchronization between midfrontal theta and typing behavior, and that breakdown in this synchronization would explain typing errors. However, it is important to note that our hypothesis does not presuppose whether brain oscillations would become phase-aligned with typing as a rhythmic sensory information, or whether brain oscillations would cyclically orchestrate processes resulting in rhythmic typing. Our results do not clearly support our hypothesis. Indeed, phase consistency (i.e., synchronization) of keystrokes did occur significantly in all participants and at various frequencies for all participants, suggesting that there is some broadband (non-frequency-specific) synchronization between typing and

neural oscillations. Analyses showed that peak synchronization seemed to occur around 8 Hz at the group level. Since the 8-Hz component was associated with midfrontal topography, this would support midfrontal theta synchronization with typing. However, inspection of participant-specific data showed that synchronization peak frequency was participant specific and ranged from 4 to 15 Hz, thus associated with components with various topographies. Therefore, although significant synchronization mostly occurred at low frequencies (and thus around midfrontal regions) at the participant level, our results do not point to a clear specificity for the theta band and for midfrontal activity. We further reasoned that if synchronization between typing and neuronal oscillation was strong, the peak synchronization frequency and the peak typing frequency should correlate at the group level. We only observed such correlation when focusing on typing frequency determined by the KDE method.

Moreover, no significant changes in phase consistency were observed between the different experimental conditions, and between the different trial outcomes, at any frequency. This suggests that the need for cognitive control did not modulate synchronization and that erroneous trials (whether they were corrected or not) were not necessarily associated with a decrease in synchronization. The average number of errors in erroneous sentences was only two, which prevented us from investigating the effect of the number of errors on synchronization. This could have influenced error-related results. Investigating such an effect would be interesting in a subsequent study but would likely require focusing on less experienced participants.

To summarize, although our results show synchronization between typing and neuronal oscillations, it does not seem to be limited to the theta band or to midfrontal regions, and typing behavior does not seem to depend on the strength of this synchronization.

### Limitations and Perspectives

This study is the first, to our knowledge, to investigate synchronization between brain activity and typing. As a result, interpretations should remain cautious and take into account some limitations.

Our results show that there was a clear variability in typing frequency across participants, with some participants showing 4.5-Hz typing frequency whereas others went up to 7.5 Hz. Such differences could participate in enhancing variability when investigating synchronization. Focusing on a group of participants sharing more similar typing frequencies could help in diminishing interindividual typing differences. However, it is important to note that our results show that, even for similar typing frequencies, the peak synchronization frequency varied from 4 to 15 Hz. This suggests that typing frequency variability probably does not entirely account for synchronization differences. Another linked limitation is that we cannot rule out that variability in typing expertise across participants



might have influenced our results. Although including random intercept for participant in the statistical models accounts for part of interindividual variability, it is possible that brain–behavior synchronization during typing varies depending on expertise. Further work including precise typing expertise information of participants will be needed to investigate this hypothesis.

With this study, we aimed at testing a specific hypothesis on midfrontal theta synchronization with typing. Thus, we chose to use a spatial filtering technique (GED) specifically designed to enhance narrow-band frequency-specific features of the signal. Although GED successfully isolated target components, it also potentially masked broader topographical networks, which do not necessarily operate in narrow-band frequencies but rather by the means of cross-frequency dynamics. An interesting perspective of this work could be to investigate such cross-frequency networks by using participant-specific peak synchronization frequency as the low frequency timing higher frequency characteristics.

Our hypothesis was based on the fact that keyboard typing is associated with the same frequency characteristics (midfrontal oscillatory dynamics) as cognitive control. These associations are often described for perception mechanisms (Zoefel & VanRullen, 2016), and an interesting contribution of our study is in showing that a comparable association is relevant for high-level real-world behavior. However, our data do not allow us to interpret the meaning of such frequency association and synchronization between behavior and neural activity.

Another aspect to bear in mind is that this study was limited to the investigation of cortical brain–behavior synchronization. The synchronization we expected might be more prominent between typing and activity in the basal ganglia or the cerebellum, which are hardly accessible through EEG. For instance, in the case of cognitive control, theta oscillations in the subthalamic nucleus have been associated with conflict resolution and inhibition and subtends functional connectivity with midfrontal cortical regions (Zavala, Zaghoul, & Brown, 2015), making it an interesting candidate structure for synchronization with typing. Another potential structure for synchronization might be the cerebellum, which is involved in the subsecond timing of voluntary movements (Bareš et al., 2019), thus in the frequency range of keyboard typing. Finally, given the complex interplay between cortical regions, the basal ganglia and the cerebellum for goal-directed behavior, it is highly unlikely that synchronization would be exclusive to one structure. Comparing the strength of synchronization between these brain regions and typing would therefore provide interesting insights on the role of brain–behavior synchronization during keyboard typing.

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## Author Contributions

Joan Duprez: Data curation; Formal analysis; Investigation; Methodology; Project administration; Software; Validation; Visualization; Writing—Original draft; Writing—Review & editing. Mitchel Stokkermans: Data curation; Investigation; Software; Writing—Review & editing. Linda Drijvers: Conceptualization; Data curation; Investigation; Project administration; Resources; Supervision; Writing—Review & editing. Michael X. Cohen: Conceptualization; Funding acquisition; Investigation; Methodology; Project administration; Resources; Software; Supervision; Validation; Writing—Review & editing.

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## Diversity in Citation Practices

A retrospective analysis of the citations in every article published in this journal from 2010 to 2020 has revealed a persistent pattern of gender imbalance: Although the proportions of authorship teams (categorized by estimated gender identification of first author/last author) publishing in the *Journal of Cognitive Neuroscience (JoCN)* during this period were  $M(\text{an})/M = .408$ ,  $W(\text{oman})/M = .335$ ,  $M/W = .108$ , and  $W/W = .149$ , the comparable proportions for the articles that these authorship teams cited were  $M/M = .579$ ,  $W/M = .243$ ,  $M/W = .102$ , and  $W/W = .076$  (Fulvio et al., *JoCN*, 33:1, pp. 3–7). Consequently, *JoCN* encourages all authors to consider gender balance explicitly when selecting which articles to cite and gives them the opportunity to report their article's gender citation balance. The authors of this paper report its proportions of citations by gender category to be:  $M/M = .708$ ;  $W/M = .042$ ;  $M/W = .125$ ;  $W/W = .125$ .

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