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Factors involved in making post-performance judgments in mathematics problem-solving

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Abstract

Background: This study examines the impact of executive functions, affective-motivational variables related to mathematics, mathematics achievement and task characteristics on fifth and sixth graders’ calibration accuracy after completing two mathematical problems. Method: A sample of 188 students took part in the study. They were divided into two groups as function of their judgment accuracy after completing the two tasks (accurate= 79, inaccurate= 109). Differences between these groups were examined. The discriminative value of these variables to predict group membership was analyzed, as well as the effect of age, gender, and grade level. Results: The results indicated that accurate students showed better levels of executive functioning, and more positive feelings, beliefs, and motivation related to mathematics. They also spent more time on the tasks. Mathematics achievement, perceived usefulness of mathematics, and time spent on Task 1 significantly predicted group membership, classifying 71.3% of the sample correctly. Conclusions: These results support the relationship between academic achievement and calibration accuracy, suggesting the need to consider a wide range of factors when explaining performance judgments.

Keywords: achievement, calibration, elementary school, executive functions, mathematical problems.

The degree of correspondence between one’s perception of performance and actual performance is referred to as calibration (Hacker, Bol, & Keener, 2008). These judgments can be made before (predictions), or after (postdictions) task completion. Postdictions tend to be more accurate and reliable than predictions, and inform about monitoring processes during the task (Bol, Hacker, Walck, & Nunnery, 2011; Hacker et al., 2008; Sheldrake, Mujtaba, & Reiss, 2014). The present study investigates which factors contribute to making accurate post-performance judgments.

As calibration informs about the status of one’s knowledge or strategies at a cognitive level, it is involved in the regulation of further effort and strategy use (Van Loon, de Bruin, Van Gog, & Van Merriënboer, 2013). The relationship between calibration and achievement has been demonstrated in different subject areas, particularly in mathematics and problem-solving (Jacobse & Harskamp 2012; Özsoy 2012; Rinne & Mazzocco, 2014).
Factors involved in making post-performance judgments in mathematics problem-solving

Specifically, calibration accuracy may explain 16-36% of the variance in mathematics achievement (Jacobse & Hariskamp, 2012), while highly calibrated students tend to perform more successfully on problem-solving tasks because they exercise a more efficient control over their problem-solving processes (Öszöy, 2012). This is coherent with the characterization of calibration as an important metacognitive monitoring process (Stolp & Zabrucky, 2009).

However, students tend to be inaccurate in their judgments, with a tendency towards over-confidence. This tendency has been confirmed in different educational stages and adult samples, suggesting that performance judgments (and biases) may be stable overtime (Bouffard & Narciss, 2011; Finn & Metcalfe, 2014; Hacker et al., 2008). Although most research has focused on examining the effects of over-confidence, under-confidence has also been demonstrated to potentially threaten learning (Duprayat, Escribe, Huet, & Régnier, 2011; Sheldrake et al., 2014). Keeping in mind that this pattern of confidence bias may become stable, it is necessary to determine what influences student performance judgments.

One of the more consistent findings in this sense concerns the relationship between academic achievement and judgment accuracy. Specifically, high-achievers have been shown to be more calibrated, and somewhat under-confident, in comparison to low-achieving students (Bol et al., 2011; Hacker et al., 2008; Öszöy, 2012). In this same direction, task characteristics (mainly difficulty) have been widely studied, leading to mixed results. Howie and Roebers (2007) reported poor calibration only for difficult questions, while other studies suggest that students would be more over-confident when the material is difficult and under-confident when it is easy (Hacker et al., 2008).

In the context of elementary school, Boekaerts and Rozendaal (2010) examined the effects of gender, type of mathematical problem (application versus computation), and time of measurement (predictions versus postdictions) on fifth graders’ calibration accuracy. Results indicated that boys were more confident than girls judging their ability to successfully solve the problems. However, they over-estimated their performance in comparison to girls, revealing poor calibration. They also found that students performed worse on application problems, and were less calibrated in these problems, and that, surprisingly, performance judgments were more accurate before task completion (i.e., predictions).

Calibration accuracy has also been shown to be influenced by affective-motivational variables. Individual differences in self-efficacy beliefs, goal orientation, and susceptibility to social influences may account for differences in judgment accuracy, while the opposite relationship has also been demonstrated (Narciss, Koerndle, & Dresel, 2011; Sheldrake et al., 2014). Specifically, Narciss and colleagues found that fifth-grade students who made accurate judgments had greater increases in satisfaction with their performance compared to students who over-estimated their performance. Calibration accuracy may also influence students’ academic itinerary in further educational stages (Sheldrake et al., 2014). In this sense, highly-calibrated students show higher levels of enjoyment and interest in mathematics than their inaccurate peers.

However, although there is a good deal of evidence suggesting the impact of the mentioned factors on calibration, less consideration has been given to domain-general factors that may also influence students’ ability to make accurate performance judgments, such as Executive Functions (EF). Defined as “the ability to maintain an appropriate problem-solving set for attainment of a future goal” (Welsh & Pennington, 1988, p. 201), EF refers to processes such as planning, goal setting, response inhibition, impulse control, attention, self-monitoring, and cognitive flexibility. These components have been shown to serve as powerful predictors of school readiness and achievement, specifically, of mathematics (Bell & Lee, 2014; Kolkman, Hoijtink, Kroesbergen, & Leseman, 2013). To date however, research on the relationship between calibration and EF is very scarce. While Rinne and Mazzocco (2014) suggest that calibration in mathematics would be linked to working memory, as well as other EF such as response maintenance, from a theoretical perspective, the components of inhibition (or impulse control), working memory, attention, planning, and cognitive flexibility may be related to the ability to evaluate one’s performance (García, González-Pienda, Rodríguez, Álvarez-García, & Álvarez, 2014; Garner, 2009). Specifically, poor attention and inhibition may lead to paying attention to irrelevant information, making it difficult to maintain a goal in working memory. Without a clear goal in mind, one would be less able to switch strategies when necessary, and even less able to accurately evaluate one’s performance. Roebers, Cimeli, Röthlisberger and Neuenschwander (2012) argued that EF and children’s awareness of the discrepancy between task demands and their own performance would progress in parallel during elementary school years. This suggests the advisability of properly examining the links between EF and calibration from a developmental perspective in childhood.

The present study

This study analyzed the impact of three sets of variables on students’ performance judgment accuracy after completing two mathematical problems. A sample of fifth- and sixth-grade students took part in the study, divided into two groups as function of their post-performance judgment accuracy (accurate/inaccurate). The studied variables were:

1) Domain-general variables (EF): impulsivity control, activity regulation, emotional control, focus, concentration, planning, working memory, organization, and flexibility.
2) Domain-specific variables: affective-motivational components concerning mathematics (perceived usefulness of mathematics, self-efficacy beliefs, intrinsic motivation, anxiety, and enjoyment), and mathematics achievement.
3) Task-related variables: perceived difficulty and time spent on the tasks.

Two main objectives were established:

1) To analyze differences in domain-general, domain-specific or task-related variables between the groups with different calibration accuracy.
2) To delimit the predictive value of these variables on student’s group membership (Accurate or Inaccurate group).

Method

Participants

A sample of 359 fifth and sixth grade Spanish students took part in the study, of which 188 were selected for further data-
Cronbach’s alphas ranged between .84 and .87. Low perception of usefulness and low enjoyment, respectively). Students were assigned to the accurate group when post-performance judgments and actual performance matched in both tasks, and to the inaccurate group when they mismatched. Students who showed a tendency towards over-confidence: 107 (56.91%) students in Task 1 and 102 (54.25%) in Task 2, judged their performance as correct while, in fact, it was incorrect. However, only 2 and 7 students respectively (less than 1%) showed the opposite pattern. Students who were accurate in only one task were excluded from the analyses to avoid haphazard responses (n = 171).

Table 1 shows groups characteristics. No statistical significant differences between groups in age, grade level and gender distribution were found.

Sample selection was made through accessibility procedures. Students volunteered for the study, and presented informed consent from their parents. Children with a diagnosis, or severe learning disabilities, were excluded from the analyses.

**Instruments**

*Domain-general variables*

The Executive Functioning Scale for Families (EFS-F; García, Álvarez-García, Cueli, González-Castro, & Álvarez, 2013) was used to evaluate Impulsivity Control, Activity Regulation, Emotional Control, Concentration, Focus, Planning, Organization, Working Memory, and Flexibility. It consists of 27 Likert-type items. Parents report the frequency with which children show different behaviors related to EF deficits (from 1 = Never, to 5 = Always). High scores indicate deficit. Reliability was high for the scale (α = .94) and moderate for its factors (.67 to .74).

*Domain-specific variables*

The Inventory of Attitudes toward Mathematics- IAM (González-Pienda et al., 2012) was used to evaluate Perceived usefulness, Self-efficacy beliefs, Intrinsic motivation, Anxiety, and Enjoyment (20 items). Items are rated on a 5-point Likert scale (1 = strongly disagree, 5 = strongly agree). Perceived usefulness and Enjoyment are worded negatively (high scores indicate low perception of usefulness and low enjoyment, respectively). Cronbach’s alphas ranged between .84 and .87.

Achievement in Mathematics was based on student final academic grades in the subject (from 0 to 10).

**Tasks characteristics**

Evidence from perceived difficulty was obtained after task completion. Students scored the degree of difficulty of each task (1 = very easy, 5 = very difficult). Time spent on the task (expressed in seconds) was registered for each task. Two tasks were used to guarantee that students performed without time pressure. No specific time restrictions were made.

**Calibration variables**

Post-performance judgments were obtained by asking students to evaluate whether they had solved the problem successfully, once they completed each task. Actual Performance was established based on the correctness of students’ answers to the problems. Both variables were expressed in dichotomous terms (success = 1, failure = 0) given the characteristics of the problems, which required a unique response.

**Procedure**

The study was conducted according to The Helsinki Declaration of the World Medical Association (Williams, 2008). The evaluation was collectively administered during a regular class, using computers. Students and parents had previously completed the questionnaires (IAM and EFS-F, respectively). Mathematical problems were given on paper and also displayed on the computer screen, through a module enabled on Moodle platform. After solving each problem, students evaluated the degree to which they had perceived the task as easy or difficult, and made a performance judgment. The questions “how easy or difficult was the problem for you?, and “do you think you have solved the problem correctly?” were presented on the computer screen, after students indicated they had finished the problems, by clicking on a button set up for this purpose. This module allowed tracking when students started and finished each task, providing a measure of performance time.

Actual performance was established based on students’ written answers. Once evidence from post-performance judgments and actual performance was obtained, students were assigned to the groups (Accurate vs. Inaccurate). The problems were taken from the book “Problem solving and comprehension” (Whimsey & Lochhead, 1999). They were based on daily and practical situations. Teachers evaluated their adequacy prior to the evaluation.

**Table 1**

<table>
<thead>
<tr>
<th>Group 1 Characteristics</th>
<th>Group 2 Characteristics</th>
</tr>
</thead>
<tbody>
<tr>
<td>N</td>
<td>79/42.2%</td>
</tr>
<tr>
<td>Age (M/SD)</td>
<td>11.04/0.808</td>
</tr>
<tr>
<td>Female</td>
<td>41/61.9%</td>
</tr>
<tr>
<td>Male</td>
<td>38/48.1%</td>
</tr>
<tr>
<td>Fifth grade</td>
<td>29/36.7%</td>
</tr>
<tr>
<td>Sixth grade</td>
<td>50/63.3%</td>
</tr>
</tbody>
</table>

Note: Group 1 = accurate; Group 2 = Inaccurate

(M)ANCOVA’s were conducted to analyze group differences in the dependent variables. Variables distribution was examined, paying special attention to skewness and kurtosis (Table 2). Effect sizes were analyzed, using Cohen’s (1988) criterion. A small association was defined as $\eta^2 = .010$ ($d = .20$), a medium association as $\eta^2 = .059$ ($d = .50$), and a large association as $\eta^2 = .138$ ($d = .80$). A discriminant analysis was carried out to determine the significance of these variables predicting group membership, following a stepwise method. As this analysis requires equal
covariance matrices in the groups, Box’s test was conducted. Age, gender, and grade level were included in the analyses as covariates, and as possible discriminative variables. Gender and grade level were coded as dummy variables for the discriminant analysis.

Results

Group differences in the studied variables

Domain-general variables

Statistically significant differences between groups in EF were found, Wilks’ $\lambda = .901$, $F(8, 145) = 5.251$, $p = .026$, $\eta^2 = .099$. Only the covariate gender was statistically significant, Wilks’ $\lambda = .913$, $F(8, 154) = 2.094$, $p = .039$, $\eta^2 = .087$. ANCOVAs indicated small group differences in Impulsivity Control, Focus, Concentration, Working Memory, Planning, and Organization levels (Table 2). Although no important EF deficits were found (scores in each component may range from 3 to 15), inaccurate students were reported by their parent as having more difficulties in all the EF components.

Domain-specific variables

Statistically significant differences between groups were found in the affective-motivational components, Wilks’ $\lambda = .872$, $F(5, 178) = 5.242$, $p < .001$, $\eta^2 = .128$. The covariates age, gender, and grade level, were not statistically significant. Accurate students showed higher motivation, and more positive perceptions and emotions related to mathematics, than their inaccurate pairs (Table 2). ANCOVAs indicated the existence of significant differences in all these variables. Effect sizes were moderate, but considerably higher in Perceived usefulness. Accurate students also showed significantly higher levels of mathematics achievement (expressed as academic grades). This variable generated the highest effect size. The covariates were not statistically significant.

### Table 2

<table>
<thead>
<tr>
<th>Group</th>
<th>Group 2</th>
<th>Differences</th>
<th>Total</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>M(SD)</td>
<td>M(SD)</td>
<td>F</td>
</tr>
<tr>
<td>Domain-general variables (EF)</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Impulsivity control</td>
<td>5.72(2.20)</td>
<td>6.54(2.52)</td>
<td>5.37</td>
</tr>
<tr>
<td>Activity regulation</td>
<td>7.00(2.29)</td>
<td>7.29(2.71)</td>
<td>.535</td>
</tr>
<tr>
<td>Emotional Control</td>
<td>6.51(2.41)</td>
<td>6.72(2.55)</td>
<td>.299</td>
</tr>
<tr>
<td>Focus</td>
<td>6.65(2.35)</td>
<td>7.73(2.60)</td>
<td>8.452</td>
</tr>
<tr>
<td>Concentration</td>
<td>6.40(2.47)</td>
<td>7.50(2.50)</td>
<td>9.108</td>
</tr>
<tr>
<td>Working Memory</td>
<td>5.23(2.14)</td>
<td>6.22(2.33)</td>
<td>8.754</td>
</tr>
<tr>
<td>Planning</td>
<td>6.38(2.47)</td>
<td>7.68(2.77)</td>
<td>10.820</td>
</tr>
<tr>
<td>Organization</td>
<td>5.87(2.66)</td>
<td>6.61(2.47)</td>
<td>3.847</td>
</tr>
<tr>
<td>Flexibility</td>
<td>6.18(2.28)</td>
<td>6.62(2.00)</td>
<td>1.989</td>
</tr>
<tr>
<td>Domain-specific variables</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Affective-motivational components</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Perceived usefulness</td>
<td>6.43(3.39)</td>
<td>8.81(3.92)</td>
<td>17.699</td>
</tr>
<tr>
<td>Self-efficacy</td>
<td>17.90(1.87)</td>
<td>16.56(3.36)</td>
<td>10.691</td>
</tr>
<tr>
<td>Intrinsic motivation</td>
<td>15.43(3.18)</td>
<td>14.42(3.04)</td>
<td>14.42(3.13)</td>
</tr>
<tr>
<td>Anxiety</td>
<td>7.67(3.08)</td>
<td>9.40(3.81)</td>
<td>11.577</td>
</tr>
<tr>
<td>Enjoyment</td>
<td>5.89(2.86)</td>
<td>7.11(3.46)</td>
<td>6.113</td>
</tr>
<tr>
<td>Mathematics achievement</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Academic grades</td>
<td>7.56(2.40)</td>
<td>5.80(2.32)</td>
<td>24.640</td>
</tr>
<tr>
<td>Task variables</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Perceived difficulty</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Task 1</td>
<td>2.78(1.39)</td>
<td>2.70(1.65)</td>
<td>.163</td>
</tr>
<tr>
<td>Task 2</td>
<td>2.75(1.41)</td>
<td>3.06(1.27)</td>
<td>2.723</td>
</tr>
<tr>
<td>Time spent on the task</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Task 1 (sec.)</td>
<td>418.46(132.69)</td>
<td>354.49(148.79)</td>
<td>8.60</td>
</tr>
<tr>
<td>Task 2 (sec.)</td>
<td>325.12(98.05)</td>
<td>321.10(122.62)</td>
<td>.084</td>
</tr>
</tbody>
</table>

Note: Group 1 = accurate; Group 2 = Inaccurate; Sec. = seconds; EF = Executive Functions; K = Kurtosis; S = Skewness
High scores in EF indicate deficit.
Perceived Usefulness and Enjoyment are negatively worded (high scores indicate low perception of usefulness and enjoyment, respectively).
Task characteristics

Perceived difficulty did not generate any significant differences. The covariates were not significant in Task 1. Age was statistically significant in Task 2, \( F(1,182) = 4.367, p = .038, \eta^2 = .022 \). Accurate students spent more time on the tasks than their inaccurate peers. These differences were significant in Task 1. No differences due to age, gender, and grade level were found.

Discriminant value of the variables predicting group membership

Box’s test confirmed equal covariance matrices (Box’s \( M = 3.127, p = .378 \)). Table 3 shows results from the discriminant analysis. Mathematics achievement (i.e., academic grades), perceived usefulness, and time spent on Task 1 significantly predicted group membership. Although significant, the inter-group variability explained by the discriminant function (“eigenvalue”) was low. The high values of Wilks’ \( \lambda \), and the moderate canonical correlation corroborated this result.

Standardized and structure coefficients represent the correlations between the discriminant function and the variables, revealing academic grades as the most influential variable.

Function coefficients provided the resulting discriminant function. This function correctly classified 71.3% of subjects (79.9% from the inaccurate group, and 59.6% from the accurate group). Students with higher levels of mathematics achievement, who perceived mathematics as more useful, and spent more time performing the first task, were better calibrated.

Discussion

This study analyzed the impact of cognitive, affective-motivation and task-related variables on students’ post-performance judgment accuracy in two mathematical problems. It is important to note in this sense, that calibration was low in this sample, which is consistent with previous literature, and the absence of effective metacognitive strategies (Bol et al., 2011). Thus, it stands to reason that calibration accuracy improves as the time spent on the task increases. This result could also be related to previous differences in impulsivity control. Higher levels of self-control may have refrained accurate students from making impulsive (and erroneous) judgments. However, this effect was mainly found in Task 1, suggesting a greater impact of this variable when task is novel.

<table>
<thead>
<tr>
<th>Table 3</th>
</tr>
</thead>
</table>

Results of discriminant analysis, using stepwise method

<table>
<thead>
<tr>
<th>Standardized coefficients</th>
<th>Structure coefficients</th>
<th>Function coefficients</th>
<th>Wilks’ Lambda</th>
<th>( F )</th>
<th>( p )</th>
</tr>
</thead>
<tbody>
<tr>
<td>1. Academic grades</td>
<td>.716</td>
<td>.767</td>
<td>.692</td>
<td>.880</td>
<td>25.195</td>
</tr>
<tr>
<td>2. Perceived usefulness</td>
<td>.561</td>
<td>.655</td>
<td>.128</td>
<td>.837</td>
<td>17.899</td>
</tr>
<tr>
<td>3. Time spent on Task 1</td>
<td>.402</td>
<td>.463</td>
<td>.003</td>
<td>.812</td>
<td>14.136</td>
</tr>
<tr>
<td>Constant</td>
<td></td>
<td></td>
<td>-1.884</td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

Eigenvalue \( .232 \)

Wilks’ \( .812 \)

\( \chi^2 (p \leq .001) \)

Canonical correlation \( .434 \)

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Finally, the predictive value of some of the studied variables on calibration was confirmed, with mathematics achievement being the most important predictor. Students with higher levels of achievement, who perceived mathematics as more useful and spent more time performing the first task, showed higher calibration. Although significant group differences in EF were found, they did not significantly predict calibration accuracy. As stepwise method was used in the discriminant analysis, it is possible that previous differences in EF were partially explained by other variables, mainly mathematics achievement. This association has been widely reported, using both performance and observation EF measures (Bull & Lee, 2014; Presentación, Segienenthal, Pinto, Mercader, & Miranda, 2015; Thorell, Veleiro, Siu, & Mohammad, 2013).

These findings advocate examining a wide range of factors when differences in calibration accuracy need to be explained. Not surprisingly, achievement was the most important variable. Nevertheless, the impact of affective-motivational variables was also demonstrated, suggesting the need to extend their study to early educational stages, in which references on this topic remain quite scarce (Adelson & McCoach, 2011; Sheldrake et al., 2014). The potential impact of performance time on post-performance judgment accuracy was also highlighted.

However, there are some limitations in the present study. First, the use of only two mathematical tasks and the dichotomous nature of the variables must be considered. Increasing the number of tasks would allow obtaining more continuous measures, as well as exploring the influence of task characteristics at a deeper level. Additionally, different types of mathematical problems should be introduced, taking into consideration Boekaerts and Rozendaal’s (2010) findings. Second, only one measure of executive functions was used, based on parents’ reports. A more detailed evaluation of this construct, including other sorts of measures, based on students’ performance, would provide additional information about the impact of these components on calibration. Third, it must be noted that the studied variables classified 71% of the sample correctly. This indicates that additional variables may be explaining group differences, which should be properly examined. Finally, although the impact of the analyzed variables on calibration accuracy has been confirmed, the opposite pattern should be explored. Thus, further research must be conducted on this issue. A comprehensive analysis of the factors involved in making accurate judgments would set the basis for more adapted and effective interventions to improve calibration, based on students’ characteristics and needs. This is necessary, taking into consideration the effects of calibration on learning and task performance.

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References


