Serotonin (5-HT) has been hypothesized to be implicated in performance monitoring by promoting behavioral inhibition in the face of aversive events. However, it is unclear whether this is restricted to external (punishment) or includes internal (response errors) events. The aim of the current study was to test whether higher 5-HT levels instigate inhibition specifically in the face of errors, measured as post-error slowing (PES), and whether this is represented in electrophysiological correlates of error processing, namely error-related negativity (ERN) and positivity. Therefore, from a large sample of human subjects (n = 878), two extreme groups were formed regarding hypothesized high and low 5-HT transporter (5-HTT) expression based on 5-HTTLPR and two additional single nucleotide polymorphisms (rs25531, rs25532). Seventeen higher (LL) and 15 lower (SS) expressing Caucasian subjects were administered the selective serotonin reuptake inhibitor (SSRI) citalopram (10 mg) intravenously in a double-blind crossover design. We found pharmacogenetic evidence for a role of 5-HT in mediating PES: SSRI administration increased PES in both genetic groups, and SS subjects displayed higher PES. These effects were absent on post-conflict slowing. However, ERN and error positivity were unaffected by pharmacogenetic factors, but ERN was decoupled from behavioral adaptation by SSRI administration in the LL group. Thus, pharmacogenetic evidence suggests that increased 5-HT levels lead to behavioral inhibition in the context of internal aversive events, but electrophysiological correlates of performance monitoring appear unrelated to the 5-HT system. Therefore, our findings are consistent with theories suggesting that 5-HT mediates the link between aversive processing and inhibition.

Key words: serotonin; 5-HTTLPR; citalopram; EEG; performance monitoring

The error-related negativity (ERN) and error positivity are EEG potentials thought to reflect early detection and evaluation of evidence for an erroneous response by a performance monitoring network originating in the anterior midcingulate cortex (aMCC; Steinhauser and Yeung, 2010; Ullsperger et al., 2014). After errors, subjects usually slow down their responses and increase their accuracy, a phenomenon known as post-error slowing (PES; Rabbitt, 1966). Single-trial ERN amplitudes have been shown to covary with the amount of slowing induced by errors, suggesting a direct link between error detection and adaptive implementation (Debener et al., 2005). Furthermore, after response conflict, subjects also show reaction slowing, which has been termed post-conflict slowing (PCS; Ullsperger et al., 2005; Verguts et al., 2011). Although PES involves the evaluation of a subjectively aversive event, this is not the case for PCS.

It is currently unclear whether 5-HT-mediated inhibition relies on this performance monitoring network or whether it is mediated in other ways, possibly subcortically. The main approach to study effects of neuromodulatory systems on human behavior are either pharmacological or genetic association studies, both of which have yielded mixed results. Some genetic studies suggest that higher 5-HT levels caused by a polymorphism

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Behavioral/Cognitive

Serotonin Reuptake Inhibitors and Serotonin Transporter Genotype Modulate Performance Monitoring Functions But Not Their Electrophysiological Correlates

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Serotonin (5-HT) has been implicated in the modulation of diverse neurocognitive functions (Lucki, 1998), and considerable attention has been paid to its role in behavioral inhibition and processing of aversive events. It has been reported repeatedly that, in the face of expected punishment, a reduction of 5-HT levels leads to a disinhibition of behavior (Soubrie, 1986; Crockett et al., 2009) such that 5-HT facilitates avoidance of detrimental actions. Such functions are controlled by a performance monitoring network, yet neural evidence for serotoninergic involvement in these abilities is sparse.

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(5-HTTLPR) at the 5-HT transporter (5-HTT) gene (SLC6A4) covary with increased neuronal performance monitoring indices, such as the ERN (Fallgatter et al., 2004), which then may lead to increased inhibition (Holmes et al., 2010); however, data are inconsistent (Olvet et al., 2010). Conversely, pharmacological challenges within the 5-HT system are complicated by the considerable degree of individual genetic variance (Veenstra-VanderWeele et al., 2000).

Thus, well controlled pharmacogenetic studies are highly desirable to elucidate the role of 5-HT in performance monitoring functions.

The aim of the current study was to systematically examine effects of an acute selective pharmacological challenge and genetic variations of the 5-HT system on PES, PCS, and the electrophysiological correlates of error processing, including the single-trial coupling between ERN and PES. We used an acute blockade of 5-HTTVs via the intravenous application of a selective serotonin reuptake inhibitor (SSRI) in a double-blind crossover design in human subjects. The genetic background was controlled for by an extreme group approach: from a large sample of genotyped Caucasian subjects (n = 878), two groups based on estimated highest (LL; n = 17) and lowest (SS; n = 15) 5-HTT expression were selected (Hu et al., 2006). We hypothesized increased PES (Crockett et al., 2012) and not PCS in S allele carriers, because PES involves evaluation of subjectively aversive events (Bourreau and Dayan, 2011). Because SSRIs increase extracellular 5-HT akin to the hypothesized difference between S and L allele carriers (Murphy et al., 2008), we furthermore expected increased PES after drug administration. If this effect is mediated via increased performance monitoring activity in the aMCC, accompanying increases in ERN and error positivity and/or coupling of these potentials and following behavioral adaptation would be expected. In contrast, differential modulation and decoupling of error-related potentials and PCS would indicate 5-HTT-dependent routes of the error signal to aMCC-based performance monitoring and post-error motor inhibition.

Materials and Methods

Participants. Of 878 subjects that had been genotyped for 5-HTTLPR polymorphisms, homozygous S and L Caucasian subjects were invited to participate in the current study. Thirty-four subjects (23 females) were included after a second genotype analysis for rs25531 and rs25532 had been performed for LL subjects (see below). Exclusion criteria were any reported history of psychiatric or neurological disorders, drug abuse, or mental retardation. In addition we were also interested in any deviations from the Hardy–Weinberg equilibrium (p > 0.054 for violation). Genotyping was performed in a two-step sequence. First, samples were loaded onto a 1.6% agarose gel in a tris-borate-EDTA solution, run for 1 h 20 min at 170 V, and visualized by etidium bromide under UV light. Samples were visualized and genotyped by at least two independent raters.

A possible problem of studies on the 5-HTTLPR is the presence of two additional single nucleotide polymorphisms (SNPs) that modulate 5-HTT expression on top of the 5-HTTLPR repeat polymorphism. rs25531 (A > G) and rs25532 (C > T) have been identified and, when not controlled for, can lead to 5-HTT expression levels for the L allele similar to those seen for the S allele (Hu et al., 2006; Wendland et al., 2008). Therefore, to maximize hypothesized group differences in 5-HTT expression, we analyzed these two additional SNPs in all L/L allele carriers and excluded low expressing genotypes. This was done by direct sequencing of similarly obtained PCR products as mentioned above. After enzymatic purification with exonuclease I and alkaline phosphatase with both amplification primers, sequences were analyzed using SeqMan DNA-Star software (Lasergene). L/L subjects who were not homozygous for both higher expressing forms at rs25531 (A) and rs25532 (C) were excluded; thus, the LL group comprises 17 carriers of the in vitro highest expressing L/L/LAC genotype (Wendland et al., 2008), and the SS group comprises 15 homozygous carriers of the S allele (Fig. 1A).

Study procedure and data acquisition. When arriving at the laboratory, all participants completed a clinical interview with the physician in charge, filled in baseline questionnaires (Table 1), and completed visual analog scales (VASs; Bond and Lader, 1974) to assess mood changes. Thereafter, in double-blind manner, either 10 mg of citalopram (Cipramil) diluted in saline (250 ml) or saline alone were administered intravenously over a period of 30 min (Fig. 1B). The order of administration was counterbalanced. Plasma levels of intravenous citalopram have been shown to remain constant over ~4 h (Lotrich et al., 2004). On average, subjects performed the task ~85 min after completion of the infusion (range, 50–125 min), which depended on the time necessary to mount EEG caps and execute other tasks. After the infusion was completed, elastic EEG caps (EasyCap) with 60 Ag/AgCl sintered electrodes were mounted in the extended 10–20 system with impedances kept below 5 kΩ. Data were recorded continuously at a 500 Hz sampling rate with BrainAmp MR plus amplifiers (Brain Products) and analyzed offline using EEGlab 12.0 (Delorme and Makeig, 2004). Subjects were instructed to respond as fast and accurately as possible. A random intertrial interval (ITI) was used that varied between 1100 and 2000 ms. Every 100 trials, subjects received written feedback on screen.
about their performance and whether they should speed up their responses. Subjects were told to speed up their responses if the number of errors committed in the incongruent condition was below 20%. Additionally, the screen was bordered by a colored frame that encouraged subjects to speed up their responses by changing the color from green to red with a delay of 12 trials when they did not commit enough errors. Each subject completed 492 trials, of which half were incongruent. The number of switches from one required response direction to the other, the total number of required responses with the left or right hand in congruent and incongruent condition, and the trial sequence between congruent and incongruent trials were exactly counterbalanced. Before the task began, all subjects completed 21 training trials on each test day.

After each session, subjects had to indicate whether they got angry when committing an error on a scale from 1 (not at all) to 10 (very). Average scores were 6.7 ± 0.3, suggesting that subjects experienced errors as aversive, and reports were not different between genetic or drug conditions [mixed linear model (MLM), all \( p > 0.70 \)].

EEG analysis. The signal was bandpass filtered from 0.5 to 42 Hz, and epochs spanning from 2 s before to 1.5 s after response onset were extracted. Erroneous and correct trials were corrected separately for artifacts and epochs that contained deviations >5 SDs of the mean probability distribution of each condition were automatically rejected. This was done so not to confuse the sometimes very high single-trial ERP amplitudes with artifacts, and no more than 5% of the trials in each condition were removed. Epoched data were demeaned and submitted to adaptive mixture independent component analysis (Palmer et al., 2012). Independent components reflecting uniform artifacts, such as eyeblinks, were removed from the data, and baseline correction from 300 to 100 ms before response onset was applied. Average and grand-average waveforms were then calculated, and event-related potential (ERP) data were measured as described in the corresponding results.

To establish the relationship between early correlates of error processing and PES, multiple robust single-trial regression analysis was used (Cohen and Cavanagh, 2011; Fischer and Ullsperger, 2013; Ullsperger et al., 2014). The regression model used the reaction time (RT) of the following trials for all error trials as the dependent variable, whereas EEG activity and following congruency were used as predictor variables including their interaction. Here, all trials in which subjects corrected their response with another button press were excluded. Single-trial data were smoothed with a running average of 10 ms before and after each data point and calculated for each data point between −250 and 500 ms surrounding the response at electrode FCz. Robust regression coefficients were scaled by their SDs and are thus comparable across subjects. We used a cluster-based permutation test to correct for multiple comparisons in the time window of interest (0−500 ms) by randomly multiplying regression weights over subjects with \( p < 0.001 \) and no more than 5% of the trials in each condition were removed. The resulting distribution of the highest sums of \( t \) values within resulting clusters of significant effects (threshold \( \alpha = 0.025 \)) was based on iteration. These were then used to determine \( p \) values for the observed effects (Maris and Oostenveld, 2007).

Statistical analyses. MLMs were used to test main effects of factors drug, genotype, and their interactions (Gueorguieva and Krystal, 2004) if not otherwise stated. An additional factor coding the current session (first or second) was introduced to account for training effects attributable to task repetition and the previously reported possible dependence of serotonergic drug effects on order of administration (Murphy et al., 2002). Thus, our main analysis consists of an MLM for the comparison of

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**Figure 1.** Study design and task structure. A, B, Study design and time schedule of drug administration and task conduction. C, The timing of the flanker task with congruent and incongruent trials shown separately. SOA, Stimulus onset asynchrony.

**Table 1. Demographics and characteristics of the sample**

<table>
<thead>
<tr>
<th></th>
<th>LL (n = 17)</th>
<th>SS (n = 15)</th>
<th>( p ) for difference</th>
</tr>
</thead>
<tbody>
<tr>
<td>Genetics</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>S-HTTLPR</td>
<td>Homozygous long</td>
<td>Homozygous short</td>
<td></td>
</tr>
<tr>
<td>rs25531</td>
<td>A/A</td>
<td>C/C</td>
<td></td>
</tr>
<tr>
<td>rs25532</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Demographics</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Gender</td>
<td>12/5</td>
<td>10/5</td>
<td>0.920</td>
</tr>
<tr>
<td>Age</td>
<td>23.8 ± 0.5</td>
<td>23.9 ± 0.9</td>
<td>0.518</td>
</tr>
<tr>
<td>Weight</td>
<td>67.8 ± 2.1</td>
<td>65.7 ± 0.7</td>
<td>0.154</td>
</tr>
<tr>
<td>Verbal IQ</td>
<td>105.8 ± 10.1</td>
<td>110.8 ± 9.0</td>
<td>0.145</td>
</tr>
<tr>
<td>Questionnaires</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>BDI-II</td>
<td>2.85 ± 0.7</td>
<td>4.03 ± 0.8</td>
<td>0.300</td>
</tr>
<tr>
<td>EPQ-RS neuroticism</td>
<td>2.94 ± 0.61</td>
<td>4.73 ± 0.67</td>
<td>0.056</td>
</tr>
<tr>
<td>EPQ-RS psychoticism</td>
<td>2.59 ± 0.33</td>
<td>3.2 ± 0.47</td>
<td>0.288</td>
</tr>
<tr>
<td>EPQ-RS extraversion</td>
<td>9.24 ± 0.71</td>
<td>8.07 ± 0.87</td>
<td>0.302</td>
</tr>
<tr>
<td>EPQ-RS lie scale</td>
<td>2.76 ± 0.50</td>
<td>0.4 ± 0.13</td>
<td>&lt; 0.001</td>
</tr>
<tr>
<td>BIS-11 total</td>
<td>62.3 ± 2.2</td>
<td>65.4 ± 2.1</td>
<td>0.306</td>
</tr>
<tr>
<td>BIS attentional</td>
<td>15.6 ± 0.6</td>
<td>16.4 ± 0.7</td>
<td>0.410</td>
</tr>
<tr>
<td>BIS motor</td>
<td>23.5 ± 0.8</td>
<td>24.1 ± 0.7</td>
<td>0.544</td>
</tr>
<tr>
<td>BIS nonplanning</td>
<td>23.2 ± 1.4</td>
<td>25.5 ± 1.1</td>
<td>0.187</td>
</tr>
</tbody>
</table>

Group LL included only the high expressing homozygous \( L_{199} \) genotype; group SS included C and T variants at rs25532. Groups did not differ with regard to demographics, IQ, and baseline depression scores. A significant difference was observed on the lie scale of the EPQ-RS, indicating higher scores of social desirability in the LL group. Additionally, a trend toward higher neuroticism scores was seen in group SS, which is in accordance with previous studies (Schinka et al., 2004). impulsiveness measured with the BIS-11 did not show group differences. Values represent mean ± SE. BDI, Beck Depression Inventory; EPQ-RS, Eysenck Personality Questionnaire Revised Short Scale; BIS-11, Barratt Impulsiveness Scale.
PES versus PCS behaviorally, which includes factors genotype, drug, session, and type of slowing (PES or PCS). Two separate MLM analyses for the EEG data were set up, including the factors genotype, medication, session, and correctness for analysis of response-related ERP amplitudes. The single-trial coupling between ERN amplitude and PES was subsequently analyzed in an MLM with the factors genotype, drug, and session. Additional control analyses were run to ensure the presence of expected task effects, and some more exploratory results are reported as described in the corresponding results for completion. A compound symmetric covariance structure was chosen because it led to best model fits. Analyzes were calculated using SPSS version 22 (IBM), and interaction effects were further analyzed by post hoc contrasts of estimated marginal means applying Bonferroni’s correction implemented in SPSS.

Unspecific drug effects and control analyses. At the end of each session, subjects filled out questionnaires and VASs again and completed a trailmaking task. This task served as a control to exclude detrimental effects of citalopram on visuomotor coordination, and different versions were used in the first and second test sessions. Citalopram did not prolong the time needed for completion (citalopram, 59 ± 2 ms; saline, 58 ± 2 ms; MLM, $F_{(1,32)} < 1, p = 0.59$), and neither were genetic group differences observed ($F_{(1,32)} < 1, p = 0.44$).

Heart rate and blood pressure were measured immediately before the intravenous cannula was placed ($t_0$), when the infusion was completed ($t_1$), and three times thereafter ($t_2$–$t_4$). Analysis of these data showed a small but robust increase of peripheral mean arterial pressure by citalopram (average of 4 mmHg) at $t_1$ and $t_4$ (both $p < 0.001$), whereas heart rates remained unchanged, matching 5-HT’s physiologic characteristics (Veeningstra-VanderWeele et al., 2000). Additionally, no difference at baseline ($t_0$) was seen ($p = 0.93$).

VASs were used to assess differences in self-reported mood changes and examined the factors calmness, alertness, and contentedness (Bond and Lader, 1974). We then compared the differences in changes of self-reports between baseline level ($t_0$) and after the session ($t_1$) between citalopram and saline conditions in an MLM analysis. We observed a trend for decreased calmness by citalopram administration ($F_{(1,32)} = 3.30, p = 0.079$) and no effects for the other two factors ($p > 0.20$). Exploratory post hoc contrasts showed that the effect on calmness was numerically larger in the SS group (change, $-0.81 ± 0.48$ cm, $F_{(1,32)} = 2.79$, $p = 0.104$) than the LL group (change, $-0.39 ± 0.45$ cm, $F_{(1,32)} = 0.76$, $p = 0.390$), yet no significant interactions between genotype and drug were observed for any item ($p > 0.50$).

The state part of the state-trait anxiety index (STAI; Spielberger et al., 1983) was used to assess differences in anxiety. Changes induced by the drug were obtained by subtracting baseline levels before each test session from the results that attended at follow-up. MLM analysis of these scores showed a trend toward larger pre–post differences in the saline condition ($F_{(1,32)} = 3.39, p = 0.075$), which was attributable to a decrease in scores seen in the saline condition ($\Delta$STAI = $-1.4 ± 0.7$ points, $-2.6 ± 2\%$ decrease relative to baseline), whereas scores remained constant under citalopram ($\Delta$STAI = $0.3 ± 0.7$ points, $1.5 ± 2\%$ increase relative to baseline). Neither the trend-level drug effect for calmness nor STAI scores correlated with drug effects on the variables of interest of PES, ERN, and error positivity (all $p > 0.18$).

In summary, these data indicate that successful drug administration and the observed trends are in accordance with effects of acute SSRI administration that may increase restlessness or anxiety initially (Burghardt and Bauer, 2013), possibly mediated by corelease of other transmitters (Fischer et al., 2015), yet no severe side effects are seen such that unspecific drug-related confounds could be ruled out.

Results

Behavioral effects

Main analysis

We calculated PES by subtracting median RTs on correct trials after errors from those that did not follow errors (Fig. 2A). Trials that represented double errors (i.e., were preceded or followed by another error) were excluded. Additionally, we performed the same calculation for trials that followed incongruent events, which are also known to induce slowing on the following trial (Ullsperger et al., 2005; Verguts et al., 2011) and submitted both to MLM analysis with factor type of slowing (PES or PCS). This analysis revealed a main effect for factor type of slowing, indicating higher slowing after incongruent events (18 ± 2 ms) compared with errors (10 ± 2 ms). Furthermore, significant interactions for type of slowing with drug ($F_{(1,96)} = 7.8, p = 0.006$) and genotype ($F_{(1,32)} = 4.2, p = 0.043$) were seen. Post hoc tests confirmed a significant drug effect on PES in which citalopram increased slowing by $9 ± 3$ ms ($F_{(1,96)} = 9.1, p = 0.003$), whereas it led to a nonsignificant decrease of PCS ($\Delta$RT, $-3 ± 3$ ms, $F_{(1,96)} = 0.9, p = 0.35$). Subjects carrying the SS genotype showed higher PES ($15 ± 3$ ms) than subjects with LL genotype ($5 ± 3$ ms, $F_{(1,55.8)} = 5.0, p = 0.029$). No difference between genetic groups was seen for PCS ($\Delta$RT, $0.6 ± 4$ ms, $F_{(1,96)} < 1, p = 0.88$). The drug effect did not depend on the genotype (interaction of genotype × drug and genotype × drug × type of slowing, both $F < 1$ and $p > 0.5$). These results indicate that slowing induced by errors but not conflict was specifically increased by citalopram and that SS subjects with presumably higher 5-HTTLPR expression also display more slowing.

Control analyses

A confirmatory MLM analysis of overall median RTs in the task showed trends for higher RTs under citalopram (360 ± 3 ms) than the saline condition (354 ± 3 ms, $F_{(1,32)} = 3.3, p = 0.077$) and toward faster RTs for subjects with genotype LL (351 ± 4 ms) than SS (362 ± 5 ms, $F_{(1,32)} = 3.45, p = 0.072$). Neither the total number of errors (Fig. 2B) nor the number of errors on incompatible trials was modulated by genotype or drug (MLM, all $p > 0.20$). Comparable with other studies, subjects in our task were slower in the more difficult incongruent trials ($\Delta$RT, $76 ± 2$ ms, $p < 10^{-10}$). Furthermore, subjects responded faster when they made an erroneous response ($\Delta$RT, $-90 ± 2$ ms, $p < 10^{-10}$). None of these factors interacted with drug or genotype (all $p > 0.29$).

Exploratory analyses

Across individuals, higher PES (but not PCS) correlated positively with a post-error increase in accuracy (PIA; Danielmeier et al., 2011) in both the
Citalopram and 5-HTTLPR do not affect ERN, CRN, or error positivity. Electrophysiological indices of performance monitoring were not significantly different between genetic and pharmacological conditions. A, B. Drug effects within both genetic groups. C. The unmedicated genetic comparison. D, E. Scalp topographies at peak latencies for ERN and error positivity, respectively, for all genotype and drug conditions, which do not indicate differences between genotypes or medication states. Shaded plots and error bars reflect SEM.

**Figure 3.** Citalopram and 5-HTTLPR do not affect ERN, CRN, or error positivity. Electrophysiological indices of performance monitoring were not significantly different between genetic and pharmacological conditions. A, B. Drug effects within both genetic groups. C. The unmedicated genetic comparison. D, E. Scalp topographies at peak latencies for ERN and error positivity, respectively, for all genotype and drug conditions, which do not indicate differences between genotypes or medication states. Shaded plots and error bars reflect SEM.

**Table 2. ERP measurements**

<table>
<thead>
<tr>
<th>Condition</th>
<th>ERN (µV)</th>
<th>CRN (µV)</th>
<th>Error positivity (µV)</th>
</tr>
</thead>
<tbody>
<tr>
<td>SS saline</td>
<td>−9.01 ± 1.0</td>
<td>1.15 ± 0.8</td>
<td>11.66 ± 0.9</td>
</tr>
<tr>
<td>SS citalopram</td>
<td>−9.49 ± 1.3</td>
<td>1.05 ± 0.9</td>
<td>10.91 ± 0.9</td>
</tr>
<tr>
<td>LL saline</td>
<td>−8.86 ± 0.7</td>
<td>1.04 ± 0.9</td>
<td>12.58 ± 0.9</td>
</tr>
<tr>
<td>LL citalopram</td>
<td>−9.43 ± 0.8</td>
<td>1.49 ± 0.7</td>
<td>12.14 ± 1.3</td>
</tr>
</tbody>
</table>

ERN and CRN were measured as the individual average minima between 0 and 100 ms after response onset at electrode FCz (±SEM). The error positivity was quantified as the individual average minima between 100 and 350 ms after response onset at electrode Cz. Neither measure was modulated by acute SSRI administration, 5-HTTLPR genotype, or their interaction.

**Main EEG analysis**

After artifact correction, on average, 43 incompatible error trials (range, 26–52) per subject remained after double errors were excluded. ERN and correct-related negativity (CRN) amplitudes were measured as the negative peak between 0 and 100 ms at electrode FCz akin to the first study reporting an effect of 5-HTTLPR on ERN amplitudes (Fallgatter et al., 2004; Fig. 3). MLM analysis confirmed a highly significant effect for factor trial type (error, correct) on EEG amplitudes (error ∆µV, −10.5 ± 0.5, F(1,83) = 445, p < 10^-10), but neither were effects for factors drug and genotype nor their interaction observed (all F values < 0.1, p values > 0.75). Exploratory contrasts within error trials revealed no effect for drug (F(1,31) < 1, p = 0.37) or genotype (F(1,31) < 1, p = 0.96) and no interaction (F(1,31) < 1, p = 0.90; for exact values, see Table 2), and neither were significant effects observed on correct trials for factors drug (F(1,31) < 1, p = 0.42), genotype (F(1,31) < 1, p = 0.92), or their interaction (F(1,31) < 1, p = 0.35). Amplitudes of the error positivity, measured as the maximum peak between 100 and 350 ms at Cz, where amplitudes were highest, were again not modulated by genotype (F(1,31) < 1, p = 0.44) or drug (F(1,31) < 1, p = 0.29; interaction, F(1,31) < 1, p = 0.88).

Furthermore, ERP peak latencies did not differ depending on any 5-HT-related factor for ERN (p values > 0.42) and error positivity (p values > 0.65). The variance in ERN latencies was somewhat larger in the LL group in the saline (mean ± SD: LL, 61 ± 21 ms; SS, 66 ± 16 ms) and citalopram (LL, 60 ± 18 ms; SS, 64 ± 14 ms) conditions.

**Single-trial regression results**

Using multiple robust regression analysis of single-trial ERN and following post-error RT, we first sought to establish the link between EEG data and PES across all subjects by testing regression coefficients collapsed over the repeated factor (drug) against zero. This revealed negative covariations in the ERN time window (ERN peak, 34 ms; t test against 0, t(30) = −4.33, p = 0.0046 cluster-based permutation test; Fig. 4A), indicating that the more negative the ERN on a single-trial level was, the higher the following RT was; this pattern is in accordance with other studies (Debener et al., 2005; Wesseling et al., 2011). Furthermore, the opposite effect was observed for error positivity amplitudes, in which a positive covariation was observed (error positivity peak, 146 ms; t(30) = 3.66, p = 0.027, cluster-based permutation test). This suggests that ERN and error positivity both are sensitive for single-trial variation and adjust subsequent slowing, likely to provide time for adjustment. No significant interaction depending on the congruency of the following trials was found. We then established the effects of drug and genotype by submitting regression coefficients (measured as response-locked minima of regression coefficients between 0 and 100 ms) to MLM analysis. This revealed a significant drug × genotype interaction: for subjects
homzygous for the L allele, citalopram significantly decreased the covariance between ERN and following RT ($\Delta b = -0.55 \pm 0.19, F_{1,331} = 7.82, p = 0.009$) which was not observed for subjects homozygous for the S allele ($\Delta b = 0.09 \pm 0.20, F_{1,331} < 1, p = 0.638); interaction drug $\times$ genotype, $F_{1,331} = 5.22, p = 0.029$; Fig. 4B). No other main effects or interactions were significant (all $p$ values $>0.12$), and no drug or genotype effects were observed for the following error-positivity-related covariation that was revealed in the initial analysis (measured as maximum regression coefficients between 100 and 200 ms, all $p$ values $>0.14$).

### Supporting EEG analyses

To increase confidence in the reported null effect on ERN amplitudes and decrease the likelihood of a methodological problem in ERP quantification that prevented us from finding a true effect, we report here additional results that were obtained by different ways of ERP quantification, inclusion of more trials, accounting for general subjective amplitude variance, and an analysis of midfrontal theta power.

First, we used the average ERN amplitude in the 20 ms surrounding the grand-average peak (64 ms) instead of the individual minimum amplitude, which may be a more sensitive measure because it reduces group variance. However, we did not observe a drug ($F_{1,331} = 1.83, p = 0.19$), genotype ($F_{1,331} < 1, p = 0.54$), or an interaction ($F_{1,331} < 1, p = 0.72$) effect. Although none of the comparisons were significant, the pattern of ERN results resembled those on PES (LL saline, $-6.8 \pm 0.7 \mu V$ vs SS saline, $-7.6 \pm 0.8 \mu V, p = 0.472$ genetic comparison without drug; LL citalopram, $-7.6 \pm 0.7 \mu V; SS$ citalopram, $-8.1 \pm 0.8 \mu V$). Results were not changed when all errors, instead of only errors on incompatible trials, were included into the analysis.

The choice of baseline can influence results of peak- and average-amplitude measures. Furthermore, baseline amplitudes in response-locked analyses can be confounded by other factors (e.g., RT). Thus, we repeated the analysis by applying a baseline correction to the response-locked data that was derived from the prestimulus activity of each trial ($-300$ to $-100$ ms) and always fell into the response–stimulus interval. The apparently more sensitive mean ERN measure ($54 \sim 74$ ms) showed no effect of drug ($F_{1,331} < 1, p = 0.44$), genotype ($F_{1,331} < 1, p = 0.86$), or their interaction ($F_{1,331} < 1, p = 0.93$). Finally, we also normalized measured ERN amplitudes by the average root mean square baseline prestimulus power, which thus should account for possible, especially morphological, intersubject differences. However, neither for minimum ($p$ values $>0.16$) nor mean ($p$ values $>0.2$) ERN amplitudes were 5-HT-related effects observed.

We also investigated genetic and pharmacological effects on midfrontal theta power. Therefore, we applied current source density transformation (Kayser and Tenke, 2006) to the EEG signal, which was then convolved with a series of complex Morlet wavelets between 3 and 40 Hz in 37 linear steps using a wavelet width of 4.5 cycles. Prestimulus ($-300$ to $-100$ ms) baseline power was subtracted from each response-locked epoch, and data were thereafter scaled by the average power in the baseline range to reflect multiples of baseline activity. Because power can only be positive and thus noise likely skews results to be larger when trial numbers are lower, the trial numbers between all subjects and across erroneous and correct responses (trials chosen as described above) were reduced to the number of the lowest category (26 trials) by randomly discarding the other trials. Theta power was then derived by collapsing power from 4 to 8 Hz and measured as the average power from $-100$ ms until $300$ ms around the response at electrodes FCz and Cz. At both electrodes and within all genotype and drug conditions, theta power was significantly larger on error (FCz: average, $616 \pm 25\%$ of baseline) compared with correct trials (FCz: average, $284 \pm 19\%$ of baseline, all $p < 10^{-6}$). However, we observed neither a genetic or pharmacological nor an interaction effect on midfrontal theta power (all $p$ values for both electrodes $>0.1$).

### Session effects

We found that overall RTs were shorter when subjects performed the task for the second time ($F_{1,32} = 9.90, p = 0.004$). Furthermore, PES decreased from 16 ms in the first session to only 4 ms in the second session ($F_{1,32} = 12.76, p = 0.001$). No interactions involving the factors genotype or drug were significant (all $p$ values $>0.19$). The number of errors was not modulated by factor session ($F_{1,32} < 1, p = 0.82$, average $n$ errors for session 1 = 63.1 ± 1.9 and for session 2 = 63.5 ± 1.9). Also, PCS was not

### Figure 4

Relationship between ERN, error positivity, and PES via multiple robust regression. A. During the time window of the ERN, a significant negative covariation between single-trial EEG in error trials and their respective following RT was observed, indicating more slowing when ERN amplitudes were higher (cluster-based permutation test, $p = 0.0046$). The opposite pattern is seen in the time window of the error positivity ($p = 0.027$) in which slowing increases when amplitudes become more positive. Data here are collapsed over drug conditions, and gray areas mark significant time windows with $p < 0.01$ uncorrected. B, A significant decrease in the covariation between single-trial ERN and consecutive slowing was seen for LL ($p = 0.009$) but not SS ($p = 0.638$) subjects when comparing saline and drug condition. Shaded plots and error bars reflect SEM.
significantly reduced in the second session (first session, 20 ± 2 ms; second session, 16 ± 2 ms; $F_{(1,32)} = 3.33, p = 0.078$; interaction type of slowing × session, $F_{(1,96)} = 3.5, p = 0.065$). This indicates that subjects increased their performance when they were more acquainted with the task indicated by lower RTs and constant error rates and also seemed to be able to decrease PES but less so PCS. This highlights the importance of accounting for session effects in repeated-measures designs, even for tasks that do not involve learning of task structures. For EGG measures, error positivity ($-1.65 ± 0.71 \mu V, F_{(1,31)} = 10.3, p = 0.003$) and CRN ($1.29 ± 0.49 \mu V, F_{(1,31)} = 10.6, p = 0.003$) amplitudes were reduced when the task was performed the second time. No effect of task repetition was observed for ERN amplitudes ($F_{(1,31)} < 1, p = 0.61$). In accordance with the observed PES reduction in the second session, although no ERN reduction was observed, we also found a session effect in that covariation between EGG and RT was reduced when the task was repeated ($\Delta b = -0.40 ± 0.14, F_{(1,31)} = 7.81, p = 0.009$).

Discussion
The aim of this study was to investigate the relationship and specificity of 5-HT to error-related aversive inhibition and its electrophysiological correlates. We found pharmacological and converging genetic evidence that higher 5-HT levels are associated with increased PES, yet no ERP effect was found.

Acute intravenous low-dose SSRI administration led to increased PES across genetic groups without affecting conflict-induced slowing. SSRIs block the 5-HTT, which prevents 5-HT reuptake and leads to increased 5-HT levels in most brain regions as evidenced via microdialysis (Beyer and Cremers, 2008), although one recent study suggests that especially cortical regions may show the opposite effect (Selvaraj et al., 2012). Furthermore, SS subjects displayed higher PES compared with the LL group, which fits well to the notion that this genotype displays higher extracellular 5-HT levels attributable to lower 5-HTT expression, as has been demonstrated in mouse models of 5-HTTLPR (Mathews et al., 2004; Murphy et al., 2008). Although the link between in vitro mRNA 5-HTT expression and in vivo measures of 5-HTT binding (Jedema et al., 2010; Murthy et al., 2010) has been highly debated, many studies report differential responses to perturbations in 5-HT neurotransmission depending on 5-HTTLPR genotype (Roiser et al., 2006; Markus and Firk, 2009). This suggests that, at least in part, genetic factors influence ongoing serotonergic neurotransmission. Thus, the pharmacological effect suggests a positive relationship between 5-HT levels and behavioral adaptation after subjectively aversive events, and this effect is compatible with the observed genetic effect. However, given the difficulty in replicating previous genetic studies (NCI-NHGRI Working Group on Replication in Association Studies, 2007), the latter finding clearly requires replication in an independent sample.

These data fit well to ideas in which 5-HT mediates behavioral inhibition only in the context of aversive events, likely to prevent the repetition of maladaptive behavior (Dayan and Huys, 2009; Boureau and Dayan, 2011; Cools et al., 2011), whereas it has been noted that serotonergic influences on inhibition are usually minimal when measured in unrewarded contexts in both humans and animals alike (Winstanley, 2011). However, there is an ongoing discussion as to how 5-HT achieves this effect. Some studies suggested that 5-HT mediates the subjective effect of aversive events. This idea is mainly based on genetic evidence in which the 5-HTTLPR S allele has been identified as a risk factor for depression after aversive life events (Caspi et al., 2003) that could be mediated by an overactive aMCC-dependent performance monitoring system (Holmes et al., 2010; Drabant et al., 2012). Supporting this idea, one study by Fallgatter et al. (2004) reported increased ERN amplitudes for S allele carriers, yet a replication attempt failed (Olvet et al., 2010). Therefore, evidence for a link between the serotonergic system and EEG correlates of error-related functions is weak.

Another suggestion is that 5-HT does not mediate the actual effect of aversive events or general behavioral inhibition but rather the link between them both (Dayan and Huys, 2008; Crockett et al., 2009). Crockett et al. (2012) showed that acute dietary tryptophan depletion (ATD), a method that lowers serotonergic neurotransmission, abolished response slowing in the context of punishment predicting stimuli in a reinforced categorization task. ATD also reduced the inhibitory effects of aversive pavlovian cues on instrumental choices and impaired the ability to passively avoid unfavorable stimuli by increasing active responding (Geurts et al., 2013). Furthermore, optogenetic stimulation of serotonergic neurons in mice increased the ability to withhold responding to obtain a primary reward without affecting locomotion (Miyazaki et al., 2014). Additionally, studies investigating serotonergic effects on instrumental choices suggest increased lose-switch behavior associated with the L allele (den Ouden et al., 2013) and increased resilience to negative after high SSRI doses in rats (Bari et al., 2010). These findings appear in line with the observed increase in PES reported here because less deliberation time may lead to less optimal choices. Furthermore, the absence of 5-HT effects on electrophysiological indices of error detection are well compatible with the notion that 5-HT mediates the link between aversive events and inhibition, because these ERPs themselves likely do not reflect this transition into behavioral adaptation.

It should be noted that, although human studies mostly found evidence for a role of 5-HT in promoting inhibition specifically when punishment was expected (Crockett et al., 2012; Geurts et al., 2013) or had to be predicted (Cools et al., 2008; Robinson et al., 2011), electrophysiological (Miyazaki et al., 2012b) and recent optogenetic evidence from rodent studies (Miyazaki et al., 2014) indicate that 5-HT may be involved in inhibition to avoid punishment but also to obtain a reward. Furthermore, interfering with serotonergic signaling of optogenetically targeted raphe neurons decreased behavior that required maintenance of motivation (Liu et al., 2014). Thus, another possible explanation for the main physiologic role of 5-HT may be to increase patience—before inhibiting or activating to avoid punishment or achieve a reward (Miyazaki et al., 2012a). Our data are compatible with this proposal of serotonergic functioning because it could be interpreted as an increase in patience to avoid repetition of a mistake accompanying increased 5-HT levels. However, more experiments that specifically orthogonalize aversive and appetitive factors with inhibition and maintenance of activity are highly desirable.

The behavioral effects on PES emerged in the absence of differences in subjective estimates of self-reported affect in response to errors, compatible with other studies (Hariri et al., 2002; Holmes et al., 2010) and in accordance with absence of effects on ERN and error positivity, which are assumed to be sensitive toward the functional significance or salience of an error (Falkenstein et al., 2000). This further suggests that 5-HT mediates inhibitory effects independent of subjective appreciation of errors and possibly the cortical performance monitoring network. Certainly the current study is not sufficiently powerful to finally rule out an association between ERN or error positivity and
5-HT, which may be found in larger samples. However, it should be noted that, even when very sensitive measures of ERN amplitudes were post hoc explored, neither drug nor genetic effects approached trend levels. Especially the absence of a drug effect in a powerful repeated-measures design with well controlled intravenous drug application and control for the genetic background suggests independence of electrophysiological indices of error processing from the 5-HT system in accordance with other studies using orally applied SSRIs (de Bruijn et al., 2006; Barnes et al., 2013). Additionally, it is unlikely that the drug dose used in the current study was insufficient to induce effects. Even an oral dose of 10 mg of citilopram has been demonstrated to block >65% of midbrain 5-HTs via in vivo single-photon emission computed tomography (Klein et al., 2006), and we observed clear behavioral and cardiovascular effects. Thus, it seems that the neurogenerators giving rise to ERN and error positivity are insensitive to acute pharmacological perturbations and likely genetic variation in the serotonergic system.

One may wonder how the behavioral 5-HT effects then arise. When the link between error detection and translation into behavioral adaptation is investigated more directly, we found differential effects of the drug in both genetic groups. For LL subjects, the strength of coupling between ERN and consecutive RT adjustments is decreased by SSRI administration, but this is not the case for SS subjects. One could speculate that higher SSRI doses may lead to interactive effects for other factors as well, for example, when assuming an inverted U-shaped relationship between 5-HT and efficient functioning (Cano-Colino et al., 2014). The decoupling of behavioral adaptation suggests an independent mechanism that increases behavioral inhibition and does not affect cortical EEG correlates. One may thus speculate that the inhibition is implemented in subcortical structures, of which especially the subthalamic nucleus (STN) seems a likely structure to mediate PES (Cavanagh et al., 2014; Siegert et al., 2014). Furthermore, the STN has long been known to receive strong serotonergic projections from raphe nuclei, in which most of the brain serotonergic neurons reside, and 5-HT injections into the STN increased firing rates (Xiang et al., 2005). At least a subpopulation of likely serotonergic raphe neurons code aversive signals (Nakamura et al., 2008), which may mediate the observed slowing effect.

Finally, the genetic findings of higher PES in the SS group also highlight that, although the S allele may be a risk factor for mood disorders, under certain circumstances, greater behavioral adaptation after mistakes may also be beneficial, because it can serve to decrease the likelihood of repeating a mistake (Homberg and Lesch, 2011). In line with this, we found a robust correlation between PES and the increase of accuracy after errors, suggesting that subjects in our task were more cautious in responding in post-error trials, which effectively kept performance high (Danielmeier and Ullsperger, 2011).

In summary, the current study provides important new evidence for a role of 5-HT in mediating rapid behavioral adaptation after response errors in that higher 5-HT levels appear to increase PES as evidenced by pharmacological effects that are corroborated by converging genetic findings. However, cortical electrophysiological correlates of error processing appear unaffected by serotonergic influences, and the behavioral effect may be mediated subcortically. This provides strong evidence for a role of 5-HT as the link between aversive processing and resulting inhibitory effects. Clearly, more imaging and electrophysiological studies are needed to elucidate the neural mechanisms underlying 5-HT-mediated inhibition, especially on the level of brainstem activity and its association to aversive coding.

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