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The Effect of Interviewers’ and Respondents’ Accent and Gender on Willingness to Cooperate in Telephone Surveys

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This article presents two real-life experiments that investigate whether an interviewer’s accent and gender combined with a respondent’s accent and gender have an impact on telephone survey cooperation rates. Expectations were based on the authority and liking principles of the compliance theory. In Study 1, 12 standard-speaking interviewers (6 men, 6 women) and 12 interviewers with a regional accent (6 men, 6 women) called 1925 male and female respondents (speaking either the standard or the regional variety). In Study 2, a female interviewer who mastered the standard accent and the regional variety, called 120 respondents from the same categories as in Study 1. The expectations were not confirmed. Interviewers with authority (male, speaking standard Dutch) had no more success than interviewers with less authority (female, speaking a regional accent), and agreement of gender and accent between interviewer and respondent had no impact on the level of cooperation of the respondents. The results seem to indicate that it is not necessary for research bureaus to reject potential employees with a regional accent or with a less authoritative voice, and that they do not need to make an effort to match interviewers and respondents in characteristics such as gender and accent.

Keywords: Response Rate; Cooperation; Authority Principle; Liking Principle; Agreement Interviewer Respondent; Selection Interviewers

Introduction

Since the 1970s, research bureaus all over the world have observed serious decreases in response rates (Steeh, 1981; De Heer, 1999; De Leeuw & de Heer, 2002; Couper & De Leeuw, 2003; Curtin et al., 2005; Groves, 2011), especially in metropolitan areas (Steeh et al., 2001; Feskens et al., 2007). This is alarming in view of the possible consequences for the generalizability of the results of surveys because nonrespondents may differ systematically from respondents and consequently increase the potential for bias due to nonresponse error (Groves & Couper, 1996a). This may lead, for instance, to faulty policy decisions by politicians. Visscher (1999a, 1999b) gives an example of the detrimental effect that nonresponse may have. In a study into living conditions and political interests, nonresponse among people with a low level of education was twice as high as among people with a high level of education. The results, therefore, showed a distorted picture.

There is no simple relationship between a nonresponse rate and nonresponse bias. Nonresponse does not necessarily lead to nonresponse bias, and there is neither a minimum response rate below which a survey estimate is necessarily biased nor a response rate above which it is never biased (Curtin et al., 2000; Keeter et al., 2000; Bethlehem, 2002; Merkle & Edelman, 2002; Groves, 2006; Groves & Peytcheva, 2008; Stoop et al., 2010). Stoop (2005) states that the hunt for the last respondents is becoming increasingly difficult and costly, and may not be altogether effective, since it does not necessarily decrease selective nonresponse (Van Ingen et al., 2009) and may even increase nonresponse bias. An exit poll experiment by Merkle et al. (1998) in the US showed, for example, that a pen incentive increased response rates, but also led to higher response bias, since it only increased the response among Democratic Party voters and not among Republicans.

Problems associated with the increasing nonresponse rates exercise minds all over the world as shown by the fact that in 2011 the 21st anniversary of the International Workshop on Household Survey Nonresponse was held. This workshop has been held every year, and in 1999, culminated in the large International Conference on Survey Nonresponse (1999). A selection of the papers from this conference was published by Groves et al. (2002). Another indication is that special issues on nonresponse have been published (i.e., the Journal of Official Statistics of June 1999, 2001 and 2011, the Journal of the Royal Statistical Society: Series A (Statistics and Society) of July 2006, the Public Opinion Quarterly of October 2006) and that an avalanche of articles on reducing nonresponse have appeared: among others, Singer et al. (1999), Groves & McGonagle (2001), Feskens et al. (2007), de Leeuw et al. (2007), Snijders et al. (2007), Durrant et al. (2010), Wenemark et al. (2011). These publications focus on three aspects (Singer, 2006): a) compensating for nonresponse by imputation and/or weighting; b) measuring nonresponse bias; and c) reducing nonresponse. The research reported in this article is an attempt to contribute to the latter.

Two types of nonresponse can be distinguished: the part of the target group that is not reached (noncontacts), and those people who are reached but refuse to take part in the interview (refusals). Nonresponse in the second sense; i.e. refusals, is the topic of our studies. Singer (2006) shows that refusals contribute more to the decrease of response rates than noncontacts; she suggests that investigation of the interaction between interviewer and respondent at the initial contact might aid in gaining
usually involves the well-known name and reputation of the
organization for which he works, but also on personal aspects.
According to the authority principle, respondents are likely to
be more willing to comply with a request from an interviewer
with authority than from an interviewer with less authority.

Liking means that people are more inclined to comply with a
request from a person they like. In a survey situation, this
means a person will more likely participate when the request
to participate is made by an interviewer whom he finds sympa-
thetic. Whether this is the case depends on many factors, but
according to Groves et al. (1992), the chances that the respon-
dent will like the interviewer increase when they are more alike;
i.e., in terms of gender, age, social and regional background,
personality, or language use. They call this agreement. We find
this agreement concept also in other socio-psychological theo-
ries like Byrne’s (1971) similarity attraction hypothesis. The
results of the study of Durrant et al. (2010) confirm the effect
of the agreement concept to a certain extent: interviewers and
respondents who shared attributes, especially educational back-
ground, tended to produce higher cooperation rates. According
to the liking principle, respondents should be more willing to
comply with a request from an interviewer with whom they
share characteristics than from an interviewer with whom they
have little in common.

The decision to refuse or comply with the request to partici-
pate in cold-call telephone surveys is usually made during the
first few sentences spoken by the interviewer (Oksenberg &
Cannell, 1988; Maynard & Schaeffer, 1997; Houtkoop-Steen-
stra & Van den Bergh, 2000) and this suggests that the decision
to refuse an interview is taken along the peripheral route of
decision (Cacioppo & Petty, 1982; Van der Vaart et al., 2005).
For that reason, it is plausible that the interviewer’s voice and
the information deduced from it by the listener plays a primary
role in the listener’s evaluation of the interviewer per the au-
thority and the liking principles. Indications that the voice of
the interviewer indeed plays an important part in the decision
of the respondent to take part in a telephone survey are found in
a number of studies. Oksenberg et al. (1986) and Oksenberg &
Cannell (1988) found that interviewers rated by experts as hav-
ing relatively high pitch, great variation in pitch, loudness, fast
rate of speaking, clear pronunciation, speaking with a standard
American accent, and perceived as sounding competent and
confident had fewer refusals than interviewers with the opposite
characteristics. The studies of Van der Vaart et al. (2005) and
Groves et al. (2008) partly supported this. Their study showed
that louder sounding and higher pitched voices were indeed
correlated with lower refusal rates, but faster speaking and be-
ing judged as more confident were not. Since findings of the
effect of interviewers’ voice characteristics on refusal rates
were not consistent over models, Groves et al. (2008) argue for
further research, especially into the interaction of interviewers
and respondents. One of the lines of research they do not men-
tion, but that seems important to us in light of the liking and the
authority principle, is the role of voice characteristics of the
respondent.

Coupland & Giles (1988) apply the agreement concept from
the liking principle to voice: the more people speak in a similar
way, the more they like each other. In the context of a telephone survey, this could imply that agreement in voice might decrease the refusal rate.

There might also be a relationship between voice and the authority principle; low voices, for example, have more authority than high voices (Tielen, 1992; Van Bezooijen, 1995; Biemans, 2000). In light of the authority principle, we can expect that interviewers with voice characteristics that are associated with authority will have fewer refusals than interviewers without those voice characteristics.

An interviewer’s voice comprises four elements: accent (revealing a possible regional and/or social origin), voice height (revealing gender), voice quality (i.e., personal voice “colour”), and prosodic characteristics (speech rate, etc.). We examined the effect of accent and gender of both interviewer and respondent on the refusal rate. We chose these two factors because they are easily noticed by the respondent, and because it is relatively easy to collect information about gender and accent from refusers based on their voices.

**Expectations**

**Gender**

According to the authority principle, interviewers with more authority should have a higher response rate than interviewers with less authority. There is a very clear relation between authority and gender: men have more authority than women (Holmes & Meyerhoff 2005; Barret & Davidson 2006), and male voices have more authority than female voices (Tielen, 1992; Van Bezooijen, 1995; Biemans, 2000). Moreover, Baruffol et al. (2001) and Hansen (2007) found that male interviewers have lower refusal rates than female interviewers in telephone surveys. This leads to Expectation 1.

**Expectation 1:** Male interviewers encounter higher response rates than female interviewers.

According to the liking principle, agreement between the interviewer and the respondent would lead to a lower refusal rate. As regards gender, this implies that male interviewers who call male respondents should be more successful than male interviewers who call female respondents, and that female interviewers who call female respondents should be more successful than female interviewers who call male respondents. The study of Durrant et al. (2010) presents some support for this since there was a tendency for female respondents to be more likely to cooperate than male respondents were when the interviewer was a woman. The liking principle leads to Expectation 2.

**Expectation 2:** The response rate is higher when interviewers call a person of the same sex than when they call a person of the opposite sex.

**Accent**

We use the term accent to designate a non-standard variety of a language; for instance, a socially determined variety (sociolect), a regional variety (dialect), or an ethnic variety (ethnolect). According to the authority principle, interviewers with more authority should have a lower refusal rate than interviewers with less authority. There is a clear relationship between authority and accent. Much research shows that people who speak the standard variety are viewed as more intelligent, more self-confident, and more competent than people with a regional accent. Standard speakers generally have higher status and more authority than speakers with an accent (Cacioppo & Petty 1982; Milroy & Milroy, 1999; Heijmer & Vonk, 2002; Kraykamp, 2005; Grondelaers et al., 2010). In the context of a telephone survey, this could mean that interviewers speaking the standard variety will have more success than those with a regional accent. Oksenberg & Cannell (1988) also found this in their US study. In the Netherlands, survey agencies usually adopt this view and ask for interviewers who speak standard Dutch in their employment advertisements (Eimers & Thomas 2000: p. 28) The use of a regional variety would be perceived as less professional and would, therefore, result in a higher refusal rate. This leads to Expectation 3.

**Expectation 3:** Interviewers who speak the standard variety have a higher response rate than interviewers who have a regional accent.

According to the liking principle, agreement between interviewer and respondent leads to lower refusal rates. Apropos of accent, this implies that interviewers who use the standard variety should have more success with respondents who also speak the standard variety than with respondents who have a regional accent, and interviewers with a regional accent should be more successful with respondents who have the same accent than with respondents who speak the standard language. The positive effect of agreement between people in a negotiation situation is, for example, found by Mai & Hoffmann (2011) in their studies of language use in service selling. When there was a fit between the dialect of the salesperson and the dialect of the client, purchase intention increased. This leads to Expectations 4a and 4b.

**Expectation 4a:** The response rate is higher when interviewers who speak the standard variety of a language call respondents who speak the same standard variety than when they call respondents with a regional accent.

**Expectation 4b:** The response rate is higher when interviewers with a regional accent call equally accented respondents than when they call respondents who speak the standard variety.

**Cooperativeness Scale**

In order to gain a deeper insight into the linguistic aspects that could play a part in refusals, we analyzed almost 1000 interviewer-respondent interactions recorded by an important Dutch survey agency prior to our real-life experiments (cf. next section). We discovered that the traditional classification as applied to interview outcomes (refusal, appointment, or success) was not adequate for our research into the effects of gender and accent on respondents’ cooperativeness on initial contact between interviewer and respondent. Respondents may be cooperative without, in the end, being able to or allowed to participate because they do not meet the selection criteria. For instance, a family member of the selected respondent answers the phone and states that the selected respondent is not available. This family member is cooperative in providing the information that leads to selection, but he or she cannot take part in the survey. Such calls are then written off as noncontacts. This is not a problem for surveys, because there was indeed no contact with the person who was selected for the survey. In order, though, to be able to gain insight into the impact of gender and accent on willingness to participate in a survey, a distinction has to be made between respondents who are cooperative and respondents who are not. For this reason, we developed a scale of
cooperativeness. The reactions of respondents (approximately 1000) to requests to participate in a survey could be classified into 12 different reaction patterns (Palmen, 2001). These 12 reaction patterns were submitted to 22 communication specialists who assigned to each situation a cooperativeness score between 0 and 100. The higher the score assigned, the more cooperative the respondent was considered to be in the reaction described. The intrarater reliability of the 22 raters was in accordance with Kendall’s W 0.84 (p < .01). Figure 1 presents the means for each of these twelve reaction patterns.

Figure 1 clearly shows that some reaction patterns are very close to each other. In order to determine which patterns differed significantly, a one-way analysis of variance with a post hoc analysis (Tukey HSD) was performed with reaction pattern (1 - 12) as an independent variable and cooperativeness as a dependent variable. The results are presented in Table 1.

Table 1 shows clearly that a number of categories do not differ significantly in cooperativeness. For example, the subjects did not make a distinction in cooperativeness between a respondent who interrupts the interviewer and slams the receiver down (1); a respondent who interrupts the interviewer and terminates the call decently (2); and a respondent who lets the interviewer finish his introductory phrases, but then frankly finishes the call (3). The post hoc (Tukey HSD) analyses showed that regarding cooperativeness, the original 12 reaction patterns could be reduced to five categories. Table 2 shows these five categories (column 1), their cooperativeness with an example (columns 3 and 4), and which of the 12 reaction patterns were taken together (column 2).

The scale’s five cooperativeness categories are represented graphically in Figure 2. The numerical value assigned to each of the five categories is the mean of the initial (non-compressed) twelve categories. This informs us of the order: a respondent who is given a 3 for cooperativeness is more cooperative than a respondent who is given a 2; 2 is more cooperative than 1, etc. Figure 2 shows that the distance between the five categories are nearly similar; the lines that connect the points have almost the same slopes; and the distance between 2 and 3 is similar to the distance between 3 and 4. This is an indication that the cooperativeness scale has interval features. According to Fienberg (2007), the variable cooperativeness can, therefore, be treated as an interval variable.

Instead of classifying a request as leading to success, refusal, or an appointment (the classic tripartite system), we assigned a cooperativeness score on a scale from 1 to 5. The traditional scale's five cooperativeness categories are represented graphically in Figure 2. The numerical value assigned to each of the five categories is the mean of the initial (non-compressed) twelve categories. This informs us of the order: a respondent who is given a 3 for cooperativeness is more cooperative than a respondent who is given a 2; 2 is more cooperative than 1, etc. Figure 2 shows that the distance between the five categories are nearly similar; the lines that connect the points have almost the same slopes; and the distance between 2 and 3 is similar to the distance between 3 and 4. This is an indication that the cooperativeness scale has interval features. According to Fienberg (2007), the variable cooperativeness can, therefore, be treated as an interval variable.

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Table 2.

<table>
<thead>
<tr>
<th>New score</th>
<th>Old scores</th>
<th>Degree of cooperation</th>
<th>Result</th>
<th>Example (I = interviewer; R = respondent)</th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td>1, 2, and 3</td>
<td>None</td>
<td>Flat refusal, without a prospect of change</td>
<td>R: No, I don’t want to participate.</td>
</tr>
<tr>
<td>2</td>
<td>4 and 5</td>
<td>Minimal</td>
<td>Refusal with explanation, possibly convertible</td>
<td>R: Well, I really don’t feel like it, and those surveys always take up so much time…</td>
</tr>
<tr>
<td>3</td>
<td>6, 7, 8, and 10</td>
<td>More than minimal</td>
<td>Helpful, but with obstacles</td>
<td>I: Oh, but this is a very interesting study, and it takes only five minutes.</td>
</tr>
<tr>
<td>4</td>
<td>9 and 11</td>
<td>Less than maximal</td>
<td>Helpful</td>
<td>R: No, thanks.</td>
</tr>
<tr>
<td>5</td>
<td>12</td>
<td>Maximal</td>
<td>Full cooperation</td>
<td>I: I have to leave the house now.</td>
</tr>
</tbody>
</table>

Figure 1.
Mean of the cooperativeness scores assigned by 22 subjects to twelve reaction patterns (0 = not cooperative at all, 100 = very cooperative).

Table 1.
Differences between the twelve reaction patterns.

<table>
<thead>
<tr>
<th>Category</th>
<th>1</th>
<th>2</th>
<th>3</th>
<th>4</th>
<th>5</th>
<th>6</th>
<th>7</th>
<th>8</th>
<th>9</th>
<th>10</th>
<th>11</th>
</tr>
</thead>
<tbody>
<tr>
<td>1</td>
<td></td>
<td></td>
<td>*</td>
<td>*</td>
<td>*</td>
<td>*</td>
<td>*</td>
<td>*</td>
<td>*</td>
<td>*</td>
<td>*</td>
</tr>
<tr>
<td>2</td>
<td></td>
<td>*</td>
<td>*</td>
<td>*</td>
<td>*</td>
<td>*</td>
<td>*</td>
<td>*</td>
<td>*</td>
<td>*</td>
<td>*</td>
</tr>
<tr>
<td>3</td>
<td></td>
<td></td>
<td>*</td>
<td>*</td>
<td>*</td>
<td>*</td>
<td>*</td>
<td>*</td>
<td>*</td>
<td>*</td>
<td>*</td>
</tr>
<tr>
<td>4</td>
<td></td>
<td></td>
<td>*</td>
<td>*</td>
<td>*</td>
<td>*</td>
<td>*</td>
<td>*</td>
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<td>*</td>
<td>*</td>
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<tr>
<td>5</td>
<td></td>
<td></td>
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<td>*</td>
<td>*</td>
<td>*</td>
<td>*</td>
<td>*</td>
<td>*</td>
<td>*</td>
<td>*</td>
</tr>
</tbody>
</table>

Note: *p < .05.
classification system, 1 and 2 would be classified as refusal, 3 as an appointment, and 4 and 5 as success.

The cooperativeness scale makes different degrees of cooperativeness—after initial contact is made—more visible than the traditional classification system does, but it is definitely not better than the traditional classification scale for reporting the response rates of surveys. The cooperativeness scale should be seen as a research instrument to be used to gain insight into which characteristics of a respondent might play a part in willingness to take part in a survey after initial contact between interviewer and respondent has been made.

We defined our expectations in (3) in terms of response rates. To assess all aspects of respondents’ cooperativeness, we developed the Cooperativeness Scale to replace the traditional refusal or response rate. To compare with our expectations, the term response rate should be read as cooperativeness score.

Design and Results of Two Real-Life Experiments

To examine and verify our expectations, we carried out two studies: a real-life study in which 24 interviewers (divided equally into four categories: male/female and standard/regional accents) called 1925 respondents from the same categories (Study 1), and a matched guise study with one female interviewer who mastered both linguistic varieties; she called 120 respondents from the same categories as in the first study (Study 2).

Study 1: Real-Life Study into the Effects of Gender and Accent on Cooperativeness

In this section, we report on the experiment designed to systematically investigate the cooperativeness scores achieved between four different types of interviewers and respondents: male and female interviewers speaking either standard Dutch or a regional variety of Dutch called respondents in the same four categories. On the basis of their cooperativeness scores, we were able to determine whether certain respondent groups were more cooperative than others, and whether their cooperativeness depended on the interviewer’s accent and gender.

Accents

Two accents were investigated in this study: standard Dutch and Dutch with a Limburgian accent (a regional accent from the south of the Netherlands). The Limburgian variety was selected for a number of reasons. Just as standard Dutch, it is a clearly recognizable variety of Dutch (Grondelaers & Van Hout, 2010), and according to dialectologists it is the most deviant compared to standard Dutch (Daan & Blok, 1969; Hoppenbrouwers, 2001). It is also a variety that is viewed by its speakers and by speakers of standard Dutch as enjoyable and pleasant (Grondelaers et al., 2010; Grondelaers & Van Hout, 2010). It seemed, therefore, to be an interesting variety to test the liking principle. If respondents with a regional accent do not value their own variety positively, they will be less likely to be willing to comply with a request from an interviewer from the same area than with a request from a speaker of the standard variety.

Interviewers

The interviewers were screened by two experienced sociolinguists to determine the pleasantness of their voices and particular speech characteristics such as loudness, pitch, and rate of speaking in order to avoid that interviewers would differ with regard to voice characteristics that affect refusals (Van der Vaart et al. 2005). Only interviewers with “normal”, non-deviant voices were selected. In addition, the sociolinguists determined whether the speech of the interviewers could be qualified as one of our research categories: clearly male or clearly female, and standard Dutch or Dutch with a Limburgian accent. For the Limburgian interviewers, this meant that their speech had to be Dutch in both grammar and vocabulary, but with a clear Limburgian colour in their pronunciation. The twelve Limburgian interviewers selected had lived in Limburg for most of their lives, and their mother tongue was a Limburgian dialect. For the standard Dutch speakers, it was important that their speech did not reveal any specific regional origin; they had to speak neutral Dutch without a regional color. The accents of the interviewers with a Limburgian accent and of the interviewers with a standard Dutch accent were easy to understand for all speakers of Dutch.

The 24 selected interviewers (6 standard Dutch men, 6 standard Dutch women, 6 Limburgian men, and 6 Limburgian women) were all younger than 25 years. They represented a rather homogeneous group, and corresponded to the average survey workforce in the Netherlands. Twenty-two of them were experienced poll-takers or telemarketers. The two interviewers who did not have this experience could easily get work in these jobs, according to the sociolinguists who screened their language; their voices and way of speaking did not betray that they were inexperienced in these jobs.

Before analyzing the data any further, we established whether there were major differences between individual interviewers or whether it was possible to consider same-sex, same-accent interviewers as one homogeneous group. To that end, we compared the mean of the cooperativeness scores the interviewers obtained with each group of respondents. The means of the cooperativeness scores were very similar among interviewers within a group. One-way analyses of variance confirmed that the differences were not significant. The results of these analyses can be found in Table 3. In our subsequent analyses, therefore, we combined the interviewer results per group.

Table 3

<table>
<thead>
<tr>
<th>Interviewers: One-way analyses of variance (per respondent group).</th>
</tr>
</thead>
<tbody>
<tr>
<td>Respondents</td>
</tr>
<tr>
<td>--------------</td>
</tr>
<tr>
<td>St. Dutch male</td>
</tr>
<tr>
<td>104</td>
</tr>
<tr>
<td>102</td>
</tr>
<tr>
<td>96</td>
</tr>
<tr>
<td>St. Dutch female</td>
</tr>
<tr>
<td>132</td>
</tr>
<tr>
<td>108</td>
</tr>
<tr>
<td>119</td>
</tr>
<tr>
<td>Limb. male</td>
</tr>
<tr>
<td>117</td>
</tr>
<tr>
<td>105</td>
</tr>
<tr>
<td>109</td>
</tr>
<tr>
<td>Limb. female</td>
</tr>
<tr>
<td>159</td>
</tr>
<tr>
<td>140</td>
</tr>
<tr>
<td>128</td>
</tr>
</tbody>
</table>
Respondents and Procedure

In order to ensure that the respondents were as similar as possible in all aspects other than accent and gender, we selected four medium-sized cities (between 50,000 and 100,000 inhabitants) in the Netherlands with similar socio-economic characteristics and infrastructure based on the statistical community data of the national statistical agency. Two Limburgian cities (Roermond & Sittard) were paired with two cities from the region in which the most standard Dutch is spoken according to Duan & Blok’s (1969) and Hoppenbrouwers’ (2001) map (Rijswijk & Velsen). For each of the four cities, we selected the phone numbers of 1500 private persons by using the electronic directory. Each interviewer was provided with a mixed list of 250 phone numbers obtained from the four cities. The interviewers alternately called standard Dutch-speaking areas and Limburgian areas. When the phone was answered by a respondent, the interviewer used the following introduction (translated from Dutch): “Good evening sir/madam, this is [name of the interviewer] of BC Research. We are currently conducting a study to assess people’s opinions of certain topics relating to the Dutch language. Could I ask you a few questions on that?” Our survey agency was called “BC Research”, BC standing for Business Communication, the department responsible for the study. To investigate the impact of gender and accent on cooperativeness, we had to ask some questions of cooperative respondents. Since we did not want to ask questions without a purpose, we asked colleagues at the Radboud University to suggest questions for which they would like answers. A survey created by Smakman (2006) was ideal for this purpose as it comprised questions on the position of Dutch attitudes towards standard Dutch and regional varieties, what standard Dutch really is, and who uses that variety. We believed this topic would be equally interesting for all four categories of respondents, and would consequently avoid a nonresponse bias (Keeter et al., 2006).

The survey was carried out on Monday through Thursday nights, between 6:30 and 9:30 p.m. All interviewers worked the same days and the same hours over the survey period. Interviewers were instructed to try to persuade hesitant respondents to participate and to give them additional information, as would be done by a real survey agency. They were told not to try any further in the case of firm, radical refusals without any cues for further discussion.

For every call, the interviewers noted whether the telephone was picked up or not. When the phone was picked up, they noted whether they had spoken to a man or a woman, and the outcome of the call (success, refusal, or appointment). When the phone was not picked up, the interviewer made a note of the telephone number, and phoned that number three other times on other days. The interviewers kept track of the number of respondents in each of the four groups, and they were instructed to continue until they had spoken to at least 25 persons in each group: 25 men and 25 women from the standard Dutch area, and 25 men and 25 women from the Limburgian area. Based on Cohen (1992), we needed an α-level of .05, a power of .8, and 16 groups with 15 respondents per cell to detect a large effect size. We decided to choose a large effect size, because small effect sizes are, in our view, not relevant for practical use in telephone surveys. The interviewers were asked to conduct 25 conversations in order to leave enough margin to be able to leave out of consideration respondents with an accent other than the two varieties investigated, children, and clearly disabled persons (hard of hearing, mentally addled, or handicapped). The number of 15 was achieved for every group; when there were more than 15 respondents, the additional data were also entered into the analysis.

Since the aim of our study was to determine whether interviewer accent and gender in combination with respondent accent and gender had an impact on the cooperativeness rate in telephone surveys, respondents’ gender and accent had to be determined. All calls—which they resulted in success, refusal, or appointment—were taped and assessed for gender and accent by two experienced sociolinguists. Assessing gender was not problematic. Accent was assessed on a scale of 1 to 1:1 meaning Dutch with a very strong regional colour, and 10 meaning perfectly neutral Dutch. Respondents who were attributed 7 or higher on this scale were qualified as speakers of standard Dutch for the purposes of this study. Respondents who were attributed 5 or less, and who were identifiable as Limburgian, were assigned to the group of Limburgian respondents.

In the end, 1925 calls were examined for the impact of gender and accent on cooperativeness. According to the traditional response classification scale, the results were 589 (31%) successful interview, 150 (8%) appointment, and 1187 (62%) refusal. Since we were only interested in whether interviewer and respondent gender and/or accent had an impact on respondents’ cooperativeness, the number of non-contacts was not relevant for our study, and for that reason they were not noted.

Results Study 1

Table 4 shows the results of a $2 \times 2 \times 2 \times 2$ analysis of variance with the cooperativeness scores as the dependent variable, and the following four factors: interviewer gender (male, female), interviewer accent (standard Dutch, Limburgian), respondent gender (male, female), and respondent accent (standard Dutch, Limburgian). Table 5 presents the means and standard deviations for each of the respondent groups.

### Table 4

<table>
<thead>
<tr>
<th></th>
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</tr>
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<tbody>
<tr>
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<tr>
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<td>.19</td>
</tr>
<tr>
<td>Gender interviewer*gender respondent (Expectation 2)</td>
<td>.011</td>
<td>.91</td>
</tr>
<tr>
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<td>.09</td>
</tr>
<tr>
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<td>.25</td>
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<tr>
<td>Possible interactions</td>
<td></td>
<td></td>
</tr>
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<td>.75</td>
<td>.38</td>
</tr>
<tr>
<td>Gender interviewer<em>accent interviewer</em>gender respondent</td>
<td>.35</td>
<td>.55</td>
</tr>
<tr>
<td>Gender interviewer<em>accent interviewer</em>accent respondent</td>
<td>.32</td>
<td>.57</td>
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<tr>
<td>Gender interviewer<em>gender respondent</em>accent respondent</td>
<td>.05</td>
<td>.82</td>
</tr>
<tr>
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<td>.00</td>
<td>.98</td>
</tr>
<tr>
<td>Accent interviewer*gender respondent</td>
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<td>.23</td>
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<td>Accent interviewer<em>gender respondent</em>accent respondent</td>
<td>1.64</td>
<td>.20</td>
</tr>
</tbody>
</table>
Expectation 1, that male interviewers encounter higher cooperativeness rates than female interviewers, was not confirmed. There was no significant main effect for gender of the interviewer on the cooperativeness rate. Male interviewers (M = 2.36, SD = 1.52) had as much success as female interviewers (M = 2.45, SD = 1.58). Nor was Expectation 2 confirmed; there was no significant interaction effect between gender of the interviewer and gender of the respondent. Male interviewers had as much success with male respondents (M = 2.31, SD = 1.55) as with female respondents (M = 2.35, SD = 1.50), and female interviewers had as much success with female respondents (M = 2.44, SD = 1.58) as with male respondents (M = 2.46, SD = 1.59).

Expectation 3, that interviewers who speak the standard variety achieve a higher cooperativeness than interviewers who have a regional accent, was also not confirmed. There was no significant main effect for the accent of the interviewer on the cooperativeness. Interviewers speaking standard Dutch (M = 2.34, SD = 1.54) had as much success as interviewers speaking Limburgian (M = 2.47, SD = 1.57). There was also no significant interaction effect between the accent of the interviewer and the accent of the respondent. Interviewers speaking standard Dutch had as much success with standard Dutch speaking respondents (M = 2.31, SD = 1.55) as with Limburgian-speaking respondents (M = 2.38, SD = 1.52), and Limburgian-speaking interviewers had as much success with Limburgian-speaking respondents (M = 2.44, SD = 1.58) as with standard Dutch-speaking respondents (M = 2.51, SD = 1.56) (Expectation 4).

None of the expectations based on either the authority principle (Expectations 1 and 3) or the liking principle (Expectations 2, 4a and 4b) were confirmed. Although we had no expectations regarding interactions of interviewers’ and respondents’ gender and accent—for example, that the cooperativeness with male interviewers would be different than with female interviewers for Limburgian respondents only—we tested all possible interactions of interviewers’ and respondents’ gender and accent in the hope of gaining more insight into the factors that play a role in cooperativeness. Table 4 shows there are no significant interactions between interviewers’ and respondents’ gender and accent. Our results show clearly that interviewers’ and respondents’ gender and accent did not influence cooperativeness. Our real life study among 1925 respondents did not confirm any of our expectations; this led us to perform a second experiment.

### Study 2: Matched-Guise Replication, One Female Interviewer with Two Accents

Table 4 showed no significant differences between the interviewers in the level of success they obtained; all 24 interviewers achieved similar cooperativeness scores. One could raise the objection that 24 interviewers could not possibly present identical behavior on all levels as they are human beings, not machines. Although they were instructed to observe the instructions strictly, small deviations may have occurred. In some cases, for instance, the respondent’s argument that he or she was probably too old to participate was not always handled in the same way; e.g., the interviewer might forget to mention the counter argument, or use different phrasing. Although these deviations do not seem to have been systemic (every interviewer had some interviews presenting slight differences), they resulted in some (inevitable) variation in the data.

Also, the voices and speech characteristics of 24 interviewers were not completely equal on all levels. Although two sociolinguists assessed the interviewers’ voices prior to our study, their voices were not identical. These differences may have been minimal, or they may possibly have affected the interviewers’ scores, but did they even out across all interviewers? In order to find out whether our results were clouded by such voice and/or behaviour differences between interviewers, we did a second real-life investigation in which we used a well-tried method from sociolinguistics: the matched-guise technique (Lambert, 1967; Giles, 1973).

In the matched-guise technique, a speaker is used who masters two varieties of a language on a native speaker’s level: for example, the standard variety and a dialect variety. Two speech fragments are recorded by this perfectly bilingual person, one in the standard language and one in the dialect variety. The fragments differ only regarding language variety; apart from that, they are the same (i.e., in voice height, voice quality, and prosodic characteristics) because both fragments are uttered by the same person. The fragments are presented to listeners who indicate their appreciation of the speaker in a scaled manner in terms of intelligence, friendliness, status, etc. Since the fragments are uttered by the same speaker, differences in appreciation of the speaker of the fragments can only be attributed to the variety they speak. This matched-guise technique was found suitable to investigate whether the results of the first study could have been clouded by voice and/or behaviour differences between interviewers. If a bilingual interviewer attained similar
results as the interviewers from Study 1, that would constitute a confirmation of the findings of that study.

In regular matched-guise studies, every subject is exposed twice to the speaker, once when the speaker uses one variety, once when he uses the other (within-groups design). This was not possible in our study, since it was not possible to call the same respondent twice with the same request. For that reason, we used a between-groups design: each respondent was contacted once and heard one variety.

The female bilingual interviewer we found, according to two sociolinguists, matched the criteria applied to the standard Dutch and Limburgian interviewers in Study 1. Unfortunately, no male interviewer who met these criteria could be found. As a consequence, only the results for the female interviewers of Study 1 could be verified: Expectations 3, 4a and 4b, and partially Expectation 2.

The female bilingual interviewer contacted 120 respondents equally divided among the same four categories (30 respondents in each cell) and in the same way as was done in Study 1. In this matched-guise study, not only respondents’ accents were determined by two sociolinguists, but also the accent of the bilingual interviewer, because there was a chance that she would accidentally speak standard Dutch and indicate it as a Limburgian accent, or the other way around.

Table 6 shows the results of a 2 × 2 × 2 analysis of variance with the cooperativeness scores as the dependent variable and the factors: interviewer accent (standard Dutch, Limburgian), respondent gender (male, female), and respondent accent (standard Dutch, Limburgian). Table 7 presents the means and standard deviations for each of the respondent groups.

The part of Expectation 2 that could be tested, does the female interviewer have more success with female respondents than with male respondents, was not confirmed. The female interviewer had as much success with female respondents (M = 2.15, SD = 1.31) as with male respondents (M = 2.72, SD = 1.70).

Expectation 3 was not confirmed; when the interviewer used standard Dutch, she did not have significantly more success (M = 2.47, SD = 1.56) than when she used Limburgian (M = 2.40, SD = 1.53). There was also no significant interaction between the accent of the interviewer and the accent of the respondent; for that reason, expectations 4a and 4b were not confirmed. Agreement between interviewer and respondent in accent (M = 2.50, SD = 1.57 for Standard Dutch, and M = 2.60, SD = 1.61 for Limburgian) did not lead to more success than disagreement.

Table 6. Female interviewer’s accent and respondents’ gender and accent: results of analysis of variance.

<table>
<thead>
<tr>
<th>Study</th>
<th>M</th>
<th>SD</th>
<th>M</th>
<th>SD</th>
<th>M</th>
<th>SD</th>
<th>M</th>
<th>SD</th>
<th>M</th>
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</tr>
<tr>
<td>Male</td>
<td>1.87</td>
<td>1.53</td>
<td>2.00</td>
<td>1.45</td>
<td>2.00</td>
<td>1.57</td>
<td>2.20</td>
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<td>2.20</td>
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<td>1.40</td>
<td>1.65</td>
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<td></td>
<td></td>
</tr>
<tr>
<td>Female</td>
<td>2.00</td>
<td>1.45</td>
<td>1.85</td>
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<td>2.20</td>
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<tr>
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<td>1.53</td>
<td>1.65</td>
<td>1.40</td>
<td>1.70</td>
<td>1.20</td>
<td>1.50</td>
<td>1.53</td>
<td>1.50</td>
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<tr>
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<td></td>
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<td></td>
<td></td>
<td></td>
<td></td>
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<td>1.50</td>
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<td>2.43</td>
<td>1.50</td>
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</table>

Table 7. Mean cooperativeness scores (M) and standard deviations (SD) for all respondent groups, per interviewer group (n = 120, 1 = no cooperation, 5 = full cooperation).

<table>
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<tr>
<th>Accent</th>
<th>St. Dutch</th>
<th>St. Dutch</th>
<th>Limburgian</th>
<th>Limburgian</th>
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<td>M</td>
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<tr>
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<td>2.00</td>
<td>1.31</td>
<td>2.00</td>
<td>1.38</td>
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<td>Male</td>
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<td>1.50</td>
<td>1.60</td>
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</tr>
<tr>
<td>Female</td>
<td>1.57</td>
<td>1.65</td>
<td>1.57</td>
<td>1.65</td>
<td>1.57</td>
</tr>
</tbody>
</table>

**Conclusion and Discussion**

We reported above on two studies into the possible impact of interviewers’ and respondents’ gender and accent on willingness to cooperate in a telephone survey. In Study 1, twelve standard-speaking interviewers (6 men, 6 women) and twelve interviewers with a regional accent (6 men, 6 women) called 1925 male and female respondents (speaking either the standard or the regional variety), to request them to take part in a telephone survey on the Dutch language. There were no significant differences in cooperation between the groups: interviewers’ and respondents’ gender and accent did not affect cooperativeness. For that reason, none of the expectations were confirmed.

The results for the female interviewers were verified and confirmed in Study 2, a matched-guise study, featuring a female interviewer who mastered both the standard accent and the regional variety. She called standard-speaking and regional-speaking male and female respondents, using the two varieties alternately. She attained similar cooperativeness scores using both varieties, and when calling either men or women. In this study, there were also no significant differences between the groups; therefore, none of the expectations were confirmed here either. The number of participants per call was purposely based on a large effect size. The reason for this choice was that detecting medium or small effects on willingness to cooperate in a telephone survey would have been less relevant.

In both studies, none of the expectations based on the **authority principle** of the compliance theory were confirmed. Male interviewers were not more successful than female interviewers (Expectation 1); interviewers who spoke standard Dutch were not more successful than interviewers who had a Limburgian accent (Expectation 3); nor were the expectations based on the **liking principle** confirmed. Cooperativeness was not higher in same-sex interviews than in interviews with a person of the opposite sex (Expectation 2), and the cooper-
tiveness was not higher when interviewers called a person who
spoke the same language variety than when they called a person
who spoke another variety (Expectation 4).

Our results for the effect of gender based on the authority
principle, Expectation 1, are in line with those of Groves &
Fultz (1985), Gerritsen & Van Bezooijen (1996), and Pickery &
Loosveldt (1998), but they do not corroborate the results of
Baruffol et al. (2001), and Hansen (2007). They found that
male interviewers had higher response rates than female inter-
viewers in telephone surveys. The difference between their
results and ours could result from the fact that the interviewers
in their studies had the ability to negotiate with the respondent
about taking part in the survey, whereas the interviewers in our
study had less opportunity because we instructed them not to
try any further in the case of firm, radical refusals. According to
Hansen (2007), the higher response rates of male interviewers
in their study may be a result of men being more impertinent
than women; male interviewers may, therefore, be more likely
than female interviewers to talk respondents into taking part in
an interview, in spite of an initial refusal.

Our results for the effect of gender based on the liking princi-
ple, Expectation 2, are not in line with Durrant et al. (2010)
who found in face-to-face interviews a tendency for female
respondents to cooperate more often than male respondents if
the interviewer is a woman. They argue that their result could
be due to either the liking principle or to the potential fear of
women for a male stranger. Since the liking principle explana-
tion raises the question of why this principle does not work for
male interviewers with male respondents, the second explana-
tion is more plausible. Moreover, the fear of a male stranger
explanation can account for the difference between our results
and those of Durrant et al. (2010). Their results were based on
the response rate in face-to-face interviews, whereas ours were
based on telephone surveys. It is likely that the fear of a male
stranger is larger in face-to-face communication than in tele-
phone communication.

Our results for the effect of the language variety used by the
interviewer, Expectation 3, are also not in line with other stud-
ies. Oksenberg & Cannell (1988) found, consistent with Ex-
pectation 3, that interviewers speaking the standard American
English had more success in the US than interviewers with a
regional accent. However, Walrave (1996) and Dehue (1997)
observed precisely the opposite in the southern part of the
Dutch-speaking language area, Flanders, where a dialect is
spoken that is viewed by its speakers as enjoyable and pleasant.
Walrave observed that a slight non-standard accent inspires
confidence on the part of the respondent and leads to a higher
response than the standard language, and Dehue states that
when a call led to a successful interview, more likely than not,
the interviewer had a regional accent. In our study, however,
both the standard Dutch interviewers and the Limburgian inter-
viewers had a similar effect on cooperativeness with all re-
pondents; however, we do not know whether our results can be
extrapolated to other varieties of Dutch. Limburgian was cho-
nen because it is known as an enjoyable and pleasant accent,
and for that reason it seemed an excellent variety to test the
liking principle in telephone surveys. It could be, though, that
studies of other varieties would have been in line with the
observations of Walrave (1996) & Dehue (1997), or with the
results of Oksenberg & Cannell (1988). This calls for further
research.

Our result that agreement between interviewer and respon-
dent in language variety did not increase cooperativeness (ex-
pectation 4a and b) is not consistent with what we expected on
the basis of the study of Mai & Hoffmann (2011) who found
that agreement between salesperson and consumer in language
variety increased purchase intention. Durrant et al. (2010) also
found that agreement between interviewer and respondent in a
number of attributes increases response rate, although they did
not study agreement in language variety. The difference be-
tween their results and ours could again be due to the fact that
their results are based on face-to-face communication and ours
on telephone communication. It is plausible that the effect of
agreement is much higher in the former than in the latter. This
definitely also calls for further research.

We found no effect of the sociolinguistically important fac-
tors of gender and accent on cooperation in the initial contact
between interviewer and respondent in telephone surveys.
Apart from the reasons mentioned above, it is also possible that
this is due to the fact there is a large difference between labo-
atory and real life. The many sociolinguistic studies on which
we based our expectations were carried out under circum-
stances that were more or less controlled, and thus artificial. An
example of this kind of research is the study by Brouwer (1989)
in which subjects evaluated speakers’ accents and personality
on the basis of a tape recording of speech fragments read by
men and women speaking standard Dutch or Dutch with an
Amsterdam accent. Moreover, in these sociolinguistic studies,
the emphasis was generally on attitudes, and not on the effect
on actual behaviour. If people with a regional accent prefer a
speaker of their own variety (attitude), it does not necessarily
imply they are more likely to participate in a telephone survey
(behaviour) when requested to do so by an interviewer using
their regional variety than they would by an interviewer with an
accent different from theirs.

If the elements of daily life are eliminated, factors such as
gender and accent might play a role; however, in everyday
communicative situations, their influence seems to be overruled
by other factors. In a society known to be “sick of surveys”, the
respondents’ attitude towards telephone surveys could be a first
important interfering factor. Most respondents probably have a
pattern of response with which they react to requests to take
part in surveys, and voice characteristics of the interviewer do
not easily alter this basic attitude. Secondly, situational factors
will play a role in the decision to be cooperative or not: Does
the topic interest me? How do I experience surveys, generally?
And time factors, of course, also play an important part in the
decision to take part in a survey. Vercruysse et al. (2011)
found that respondents with less free time in Flanders more
often decline to take part in a survey. Van Ingen et al. (2009),
however, show that there is no direct proportional relationship
between business and reluctance to take part in a survey; they
found that people who are busy in work, sport, and volunteer
projects are often more inclined to take part in a survey than
people who don’t have such activities. It is not plausible that an
interviewer’s characteristics will have an impact on these situ-
ational factors. If the interviewer does not appear sympathetic,
this will probably have a negative effect, but a sympathetic
impression (for example, because of agreement in gender and/
or accent) will not automatically lead to cooperation. Our re-
results seem to indicate that various factors that were carefully
eliminated in laboratory experiments are decisive in everyday
life. For communication research, this implies that data col-
lected in authentic communicative situations constitute a valu-
able and necessary complement to data collected in laboratory situations. As the survey results are decisive, measures aimed at increasing cooperativeness should focus on the way the fieldwork is organized. It is important for research bureaus to know they probably do not have to exclude potential employees with a regional accent, or female interviewers (which could lead to a considerable increase in the number of potential interviewers), and that it is probably not necessary to make efforts to match certain characteristics of interviewers and respondents.

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